

Bootstrap for mortality projections on dependent data

V. D'Amato¹, S. Haberman², G. Piscopo³, M. Russolillo¹

*1 Department of Economics and Statistics, University of Salerno – Italy -e-mail:
vdamato@unisa.it, mrussolillo@unisa.it*

*2 Faculty of Actuarial Science and Insurance, Cass Business School, City University,
London – UK - e-mail: s.haberman@city.ac.uk*

*3 Department of Mathematics for Decisions, University of Florence, Italy - e-mail:
gabriella.piscopo@unifi.it*

AGENDA

- The motivation
- The Lee Carter Model
- The Lee Carter Sieve Bootstrap
- Numerical Applications

Motivation

Because of the nonlinear nature of the quantities of interest, such as life expectancy, annuity premiums and so on, an analytic approach to the calculation of prediction intervals is intractable, so that it is necessary to resort to a *simulation* approach.

Motivation

The presence of *dependence* across time leads to systematic over-estimation or under-estimation of uncertainty in the mortality estimates, caused by whether negative or positive dependence dominates.

Motivation

The correlation structure between the residuals has to be tackled. Otherwise prediction intervals for projections underestimate the actual longevity risk.

In other words, it is necessary to assess a significant and further source of risk: a sort of *dependency risk*.

The Lee Carter model

Lee and Carter (1992) suggested a log-bilinear form for the force of mortality:

$$m_{xt} = \exp(\alpha_x + \beta_x k_t + u_{xt})$$

$$\ln(m_{xt}) = \alpha_x + \beta_x k_t + u_{xt}$$

$$\sum_t k_t = 0 \qquad \sum_x \beta_x = 1$$

The LC Sieve Bootstrap

In the literature, there is more than one **bootstrap** method for **dependent data** as for example block, local, wild, Markov bootstrap, sub-sampling and **sieve**.

Choi and Hall (2000) show that the sieve bootstrap has substantial advantages over blocking methods, to such an extent that block-based methods are not really competitive. In particular, other authors show that the sieve bootstrap outperforms the block bootstrap (Hardle et al. 2003).

The LC Sieve bootstrap

Notation:

u_{xt} error term

ε_{xt} innovation term

r_{xt} estimated innovation or residual

\bar{r}_{xt} mean value of the residuals

$r_{xt} - \bar{r}_{xt}$ centred residuals

\hat{F}_r ecdf of residuals

u_{xt}^* bootstrap error ε_{xt}^* IID term from \hat{F}_r

The LC Sieve Bootstrap

The Scheme:

The error term is approximated by an $AR(\infty)$ representation:

$$u_{xt} = \sum_{j=1}^{\infty} \varphi_j u_{xt-j} + \varepsilon_{xt} \quad x = 1, 2, \dots, m$$

The LC Sieve Bootstrap

The Steps:

1. Fit the model and obtain the OLS estimates :

$$\hat{u}_{xt} = \sum_{j=1}^{\hat{p}(n)} \varphi_j \hat{u}_{xt-j} + \varepsilon_{xt} \quad x = 1, 2, \dots, m$$

The Lc Sieve Bootstrap

The Steps:

2. Specify the lag length $\hat{p}(n)$
by BIC, AIC, etc

The LC Sieve Bootstrap

The Steps:

- 3. Calculate the autoregressive coefficients by the Ordinary Least Squares or by using the Yule-Walker method*

$$\hat{\phi}_j, \quad j = 1, \dots, \hat{p}(n)$$

The LC Sieve Bootstrap

The Steps:

4. Calculate the residuals (or estimated innovations) associated with $\hat{\phi}_j$ according the following formula:

$$r_{xt} = \hat{u}_{xt} - \sum_{j=1}^{\hat{p}(n)} \hat{\phi}_j \hat{u}_{xt-j} \quad x = 1, 2, \dots, m$$
$$t = \hat{p}(n) + 1, \dots, n$$

The LC Sieve Bootstrap

The Steps:

5. *Calculate the centred residuals*

$$\tilde{r}_{xt} = r_{xt} - \bar{r}_{xt}$$

The LC Sieve Bootstrap

The Steps:

6. Define the empirical distribution function of the centred residuals

$$\hat{F}_{xr}(y) = \frac{1}{n-p} \sum_{t=p+1}^n 1_{\{\tilde{r}_{xt} \leq y\}}$$

The LC Sieve Bootstrap

The Steps:

7. Draw ε_{xt}^* IID terms from \hat{F}_{xr} with replacement

The LC Sieve Bootstrap

The Steps:

8. Bootstrap u_{xt}^* are simulated by recursion according to the bootstrap regression model:

$$u_{xt}^* = \sum_{j=1}^{\hat{p}(n)} \hat{\varphi}_j u_{xt-y}^* + \varepsilon_{xt}^* \quad x = 1, 2, \dots, m$$

The LC Sieve Bootstrap

Summary:

In other words, the values of ε_{xt}^* are obtained by randomly sampling with replacement from \hat{F}_{xr} and consequently the simulated u_{xt}^* are computed and the m_{xt}^* are mapped. Finally the estimates $\hat{\alpha}_x^*$, $\hat{\beta}_x^*$, $\hat{\kappa}_t^*$ are obtained by fitting the log-bilinear structure to the m_{xt}^*

Numerical Application

Application scheme:

- Model Fitting
- Analysis of residuals
- Simulation algorithm
- Comparison of the results

Numerical Application

Dataset:

The population data is composed by the Italian male from 1980 up to 2006 from 0 up to 100 years, collected from Human Mortality Database. The death rates above age 100 have been aggregated in an open age group 100+.

Numerical Application

Italy: male death rates (1980-2006)

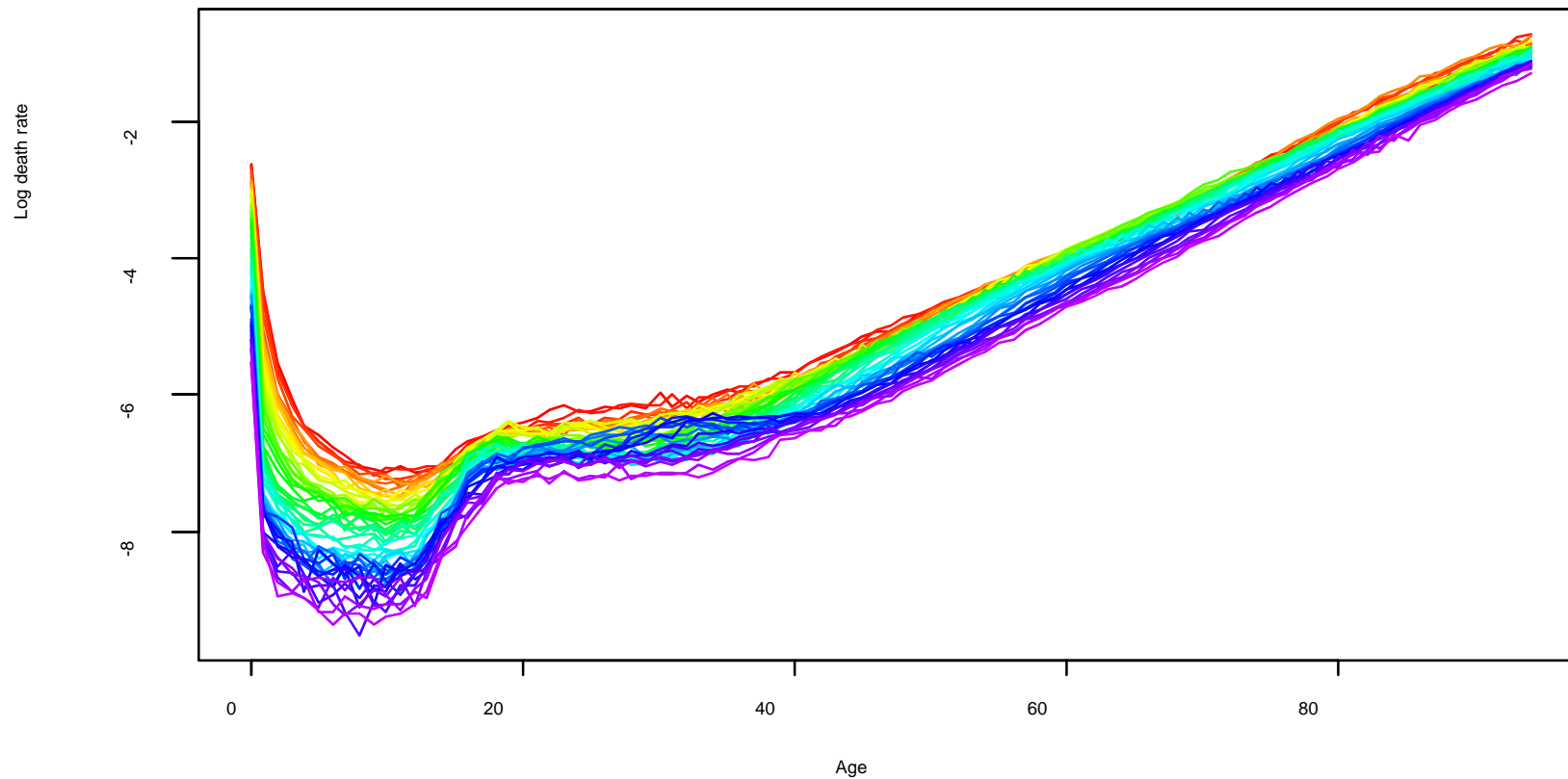


Figure 1- log death rates - Italian male population, age from 0 to 100

Numerical Application

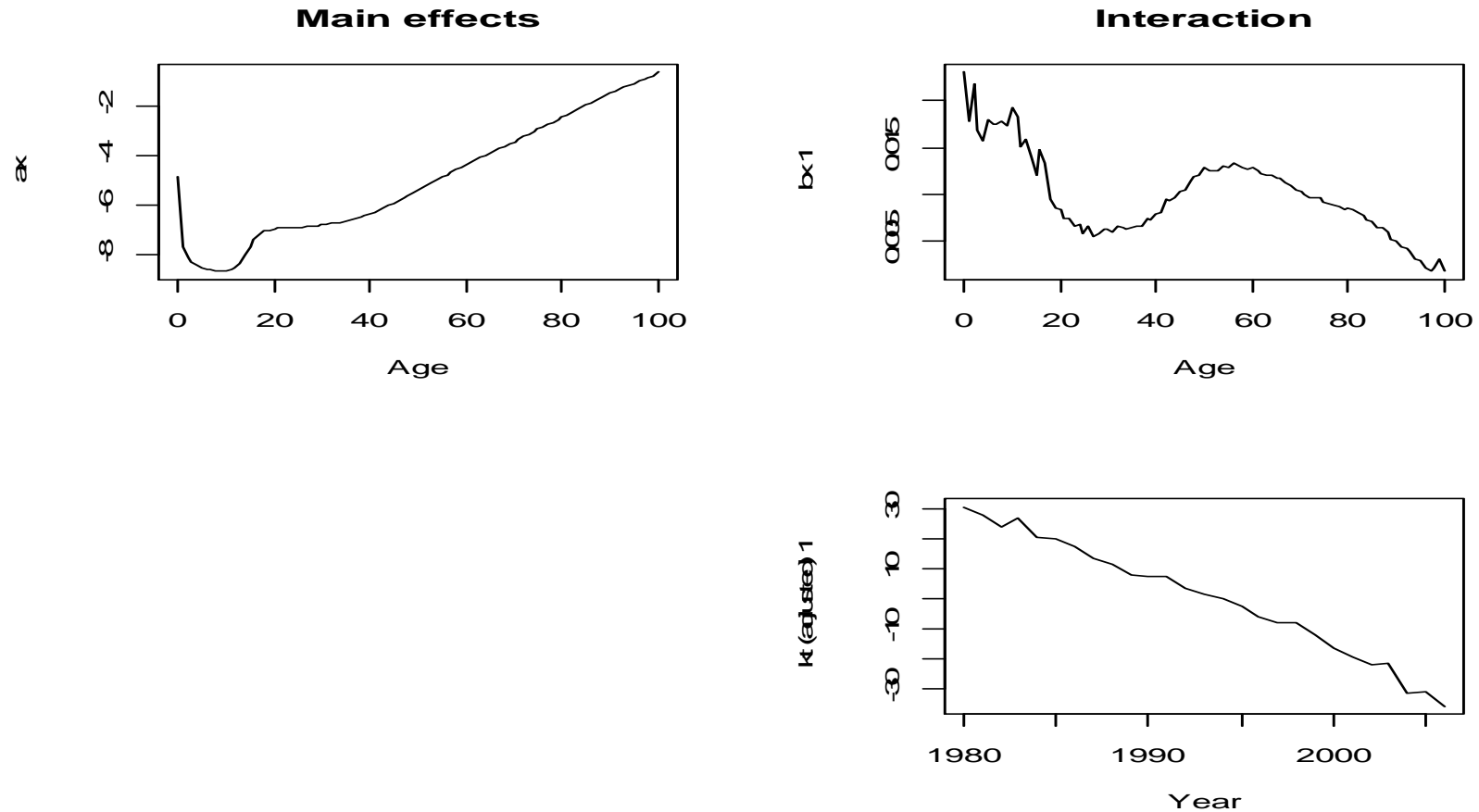


Figure 2- ax , bx , kt , basic LC model - Italian male population, age from 0 to 100

Numerical Application

ERROR MEASURES BASED ON MORTALITY RATES			
<i>Averages across ages:</i>			
ME Mean error	MSE Mean Squared Error	MPE Mean Percentage Error	MAPE Mean Absolute Percentage Error
-0.00008	0.00028	0.01102	0.08361
<i>Averages across years:</i>			
IE Integrated Error	ISE Integrated Squared Error	IPE Integrated Percentage Error	IAPE Integrated Absolute Percentage Error
-0.00455	0.01930	1.10273	8.19838

Numerical Application

ERROR MEASURES BASED ON LOG MORTALITY RATES			
Averages across ages:			
ME Mean error	MSE Mean Squared Error	MPE Mean Percentage Error	MAPE Mean Absolute Percentage Error
0.00367	0.01487	-0.01132	0.03596
Averages across years:			
IE Integrated Error	ISE Integrated Squared Error	IPE Integrated Percentage Error	IAPE Integrated Absolute Percentage Error
0.36583	1.43059	-1.00498	3.40482

Numerical Application

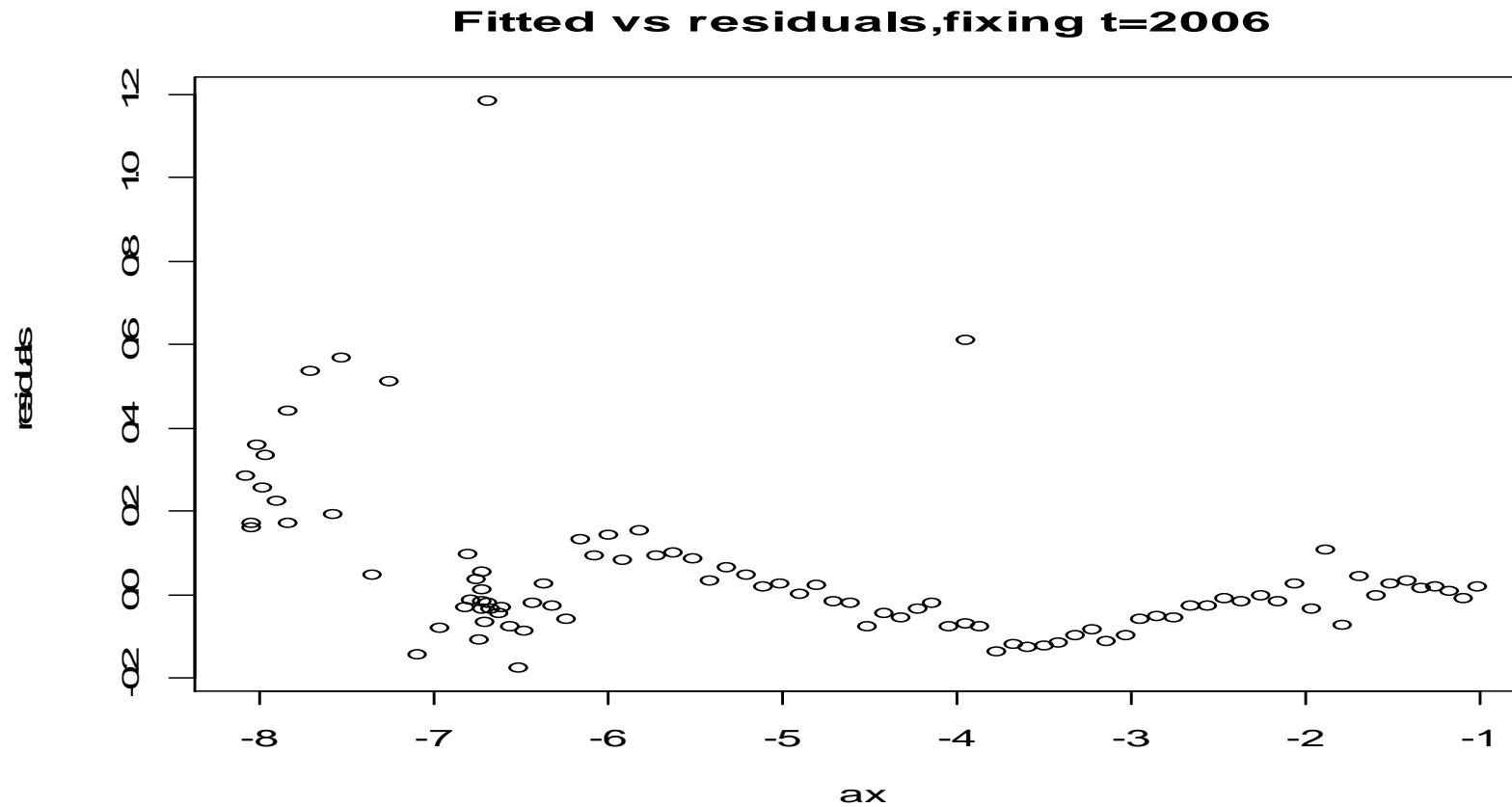


Figure 3 - Fitted ax vs residuals

Numerical Application

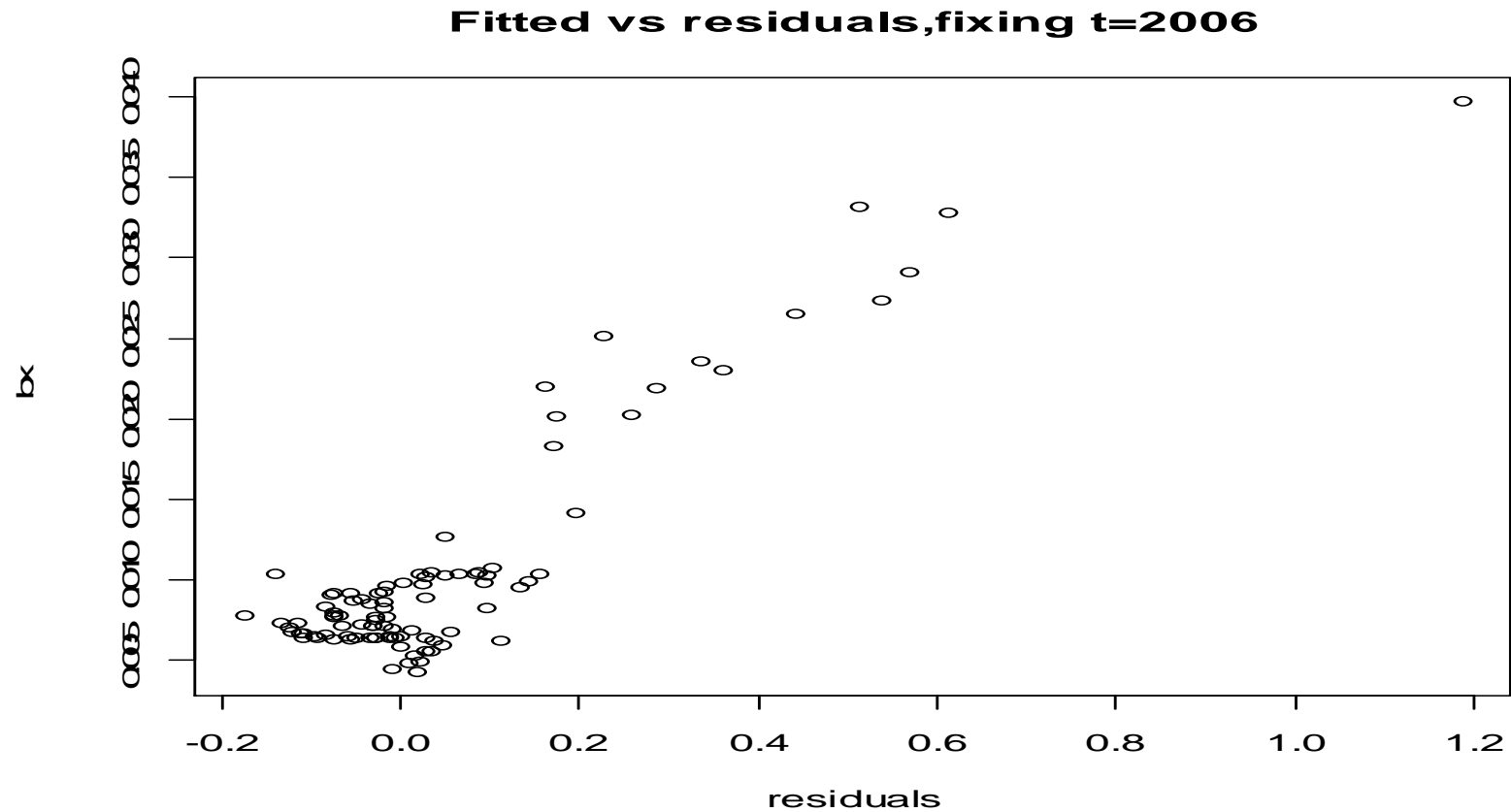


Figure 4 - Fitted bx vs residuals

Numerical Application

Fitted vs Residuals, age=65

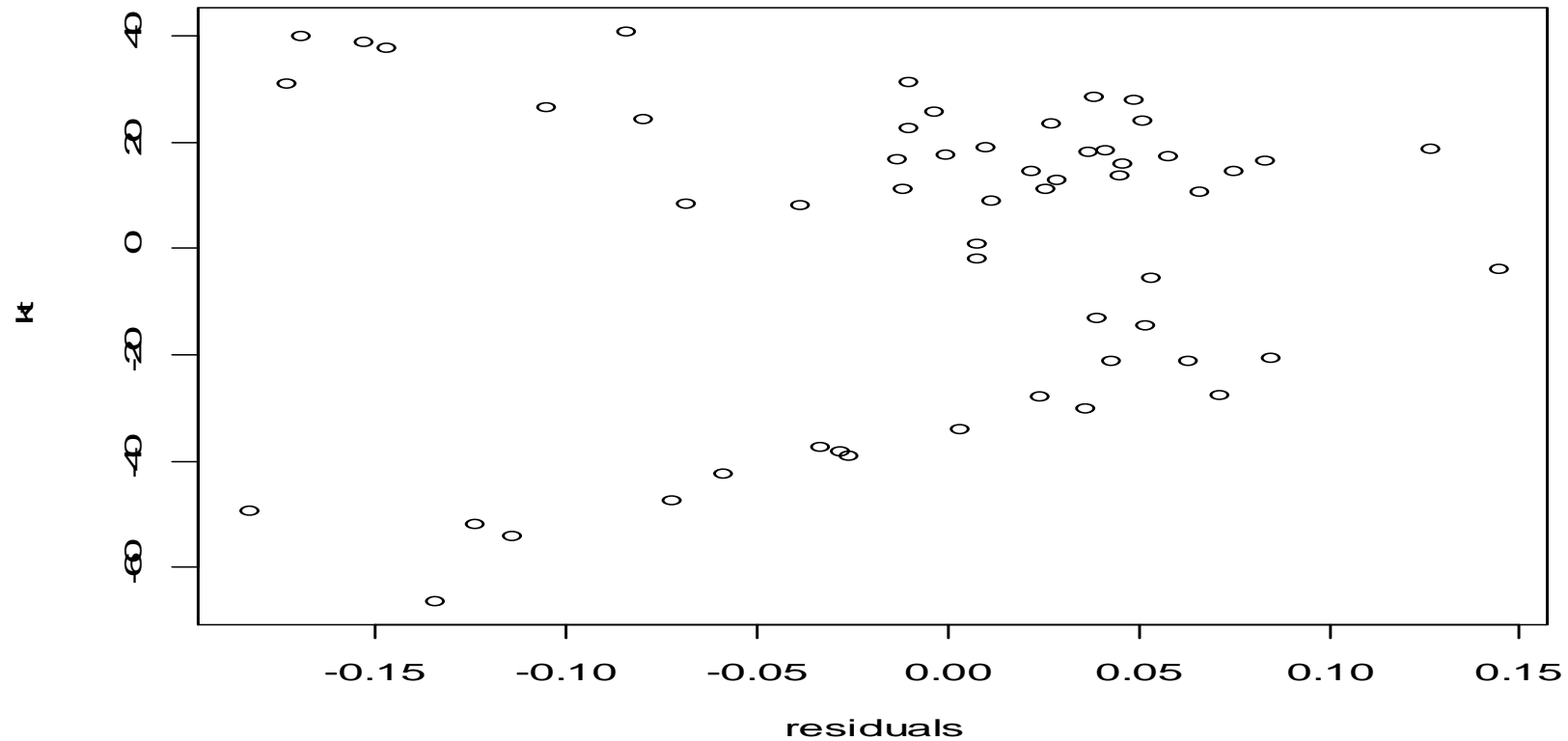


Figure 5 - Fitted kt vs residuals

Numerical Application

Residuals for Italian male, basic Lee Carter

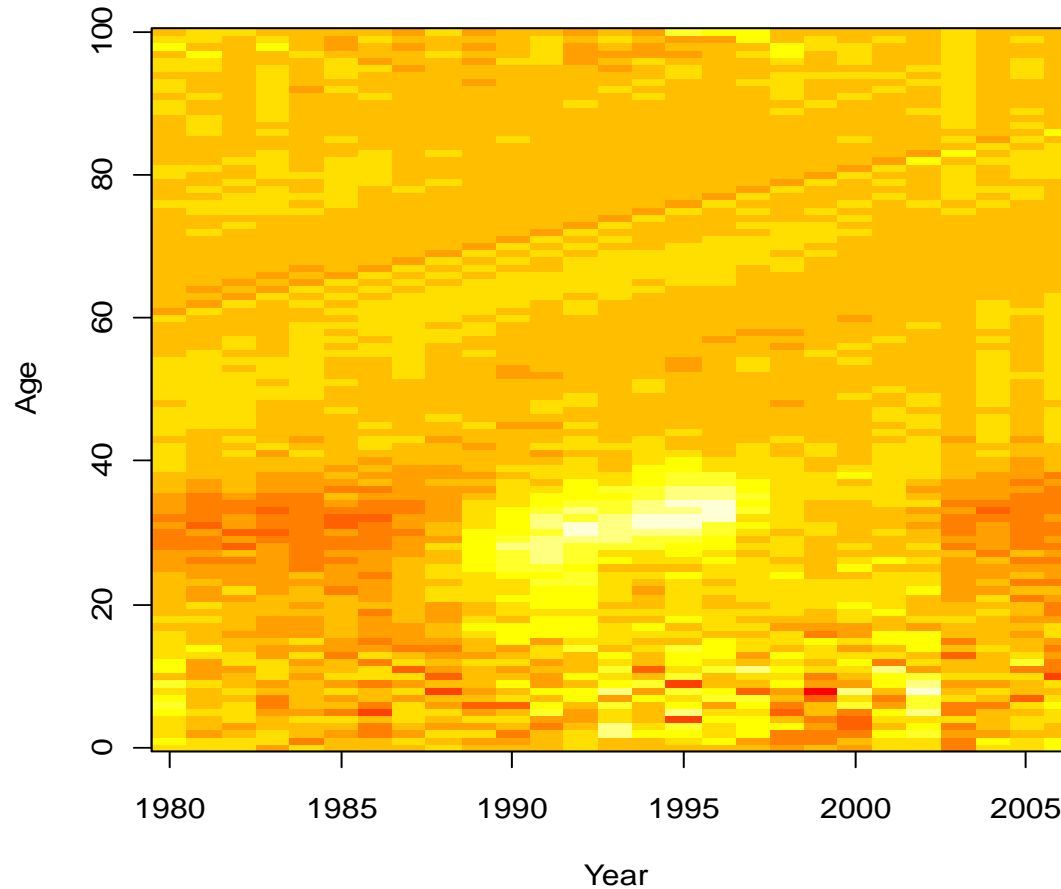


Figure 6 – Residuals years vs age – basic LC on Italian c

Numerical Application

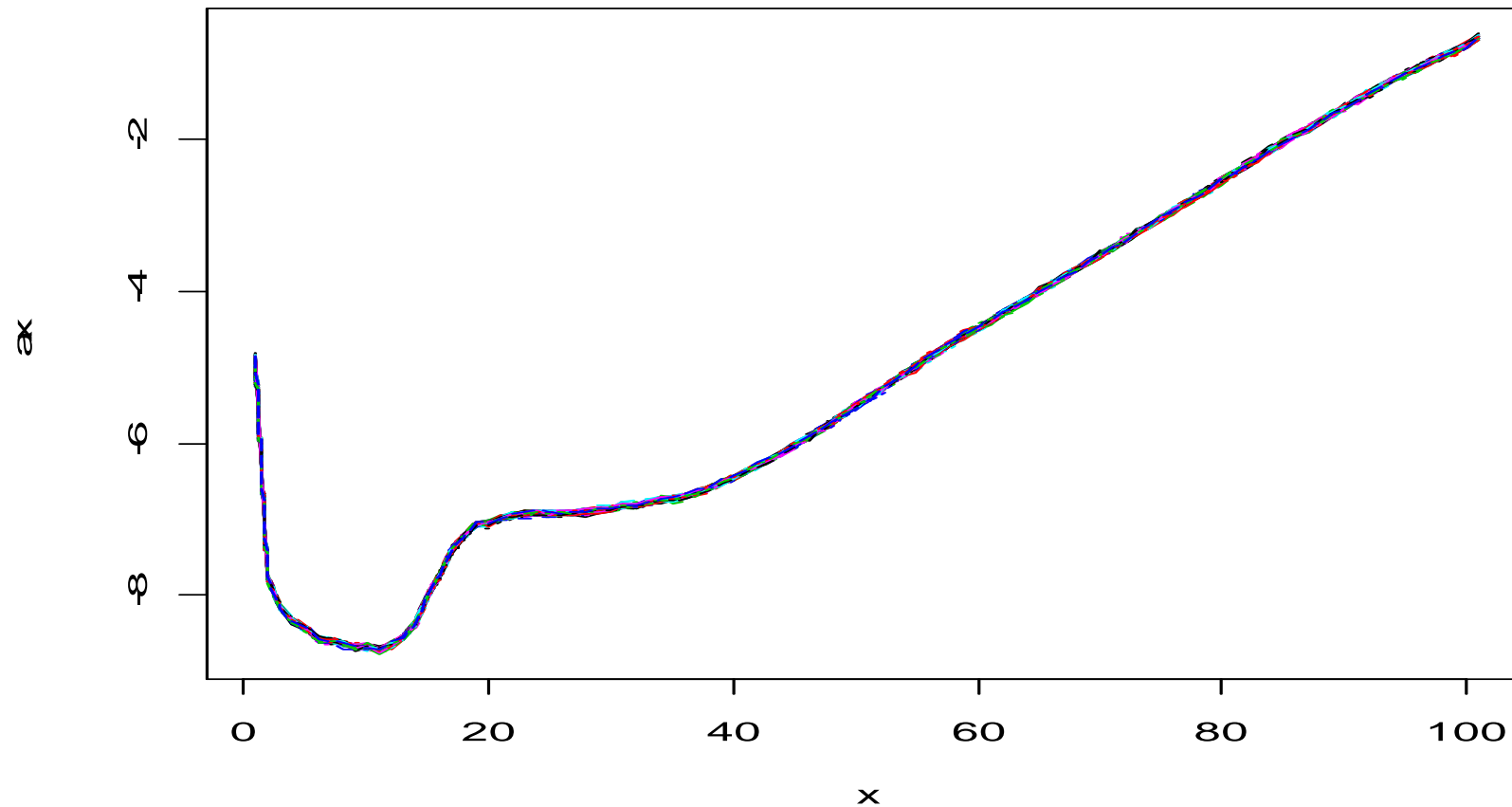


Figure 7 – Paths for ax – Sieve Bootstrap

Numerical Application

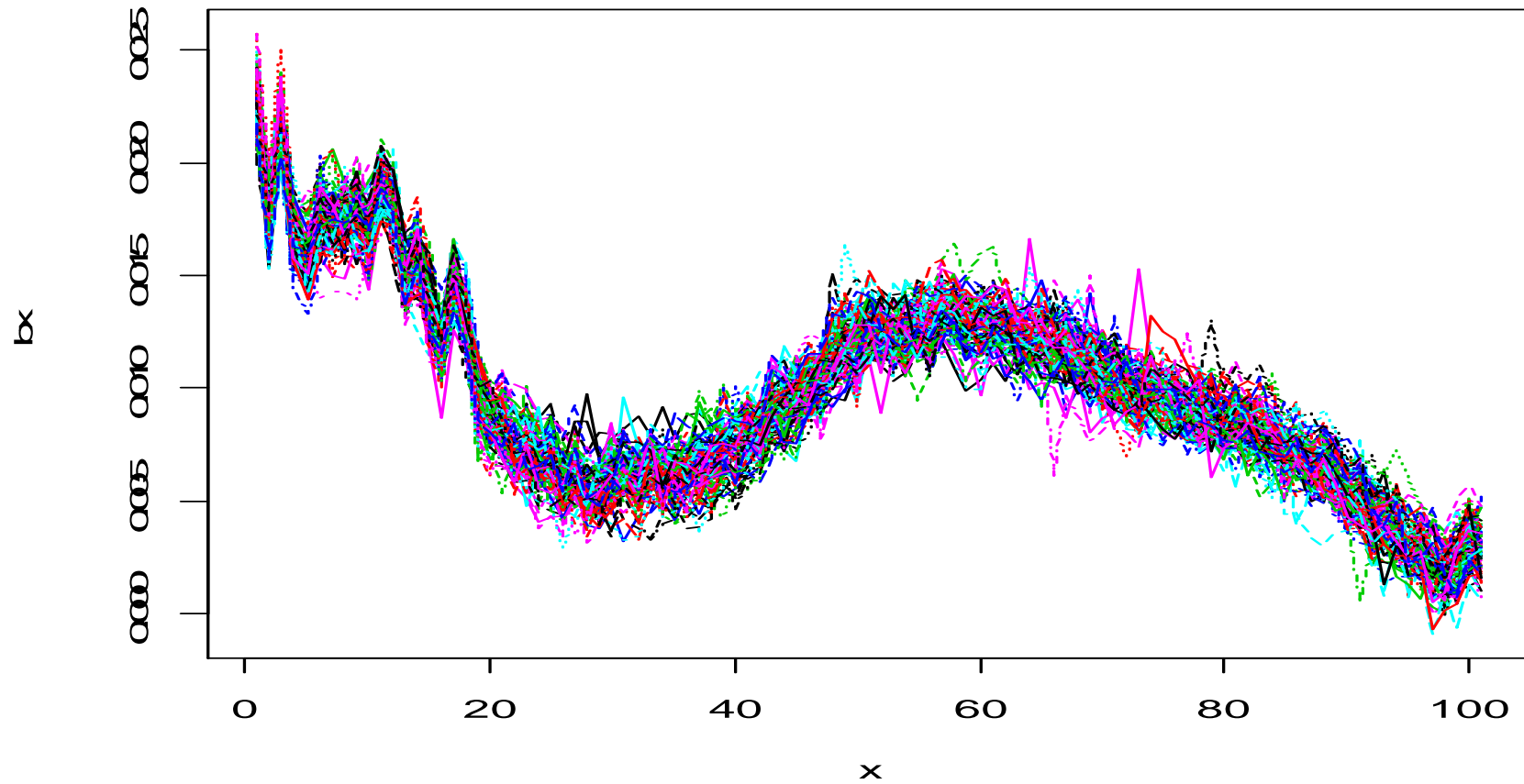


Figure 8 - Simulated paths for bx – Sieve Bootstrap

Numerical Application

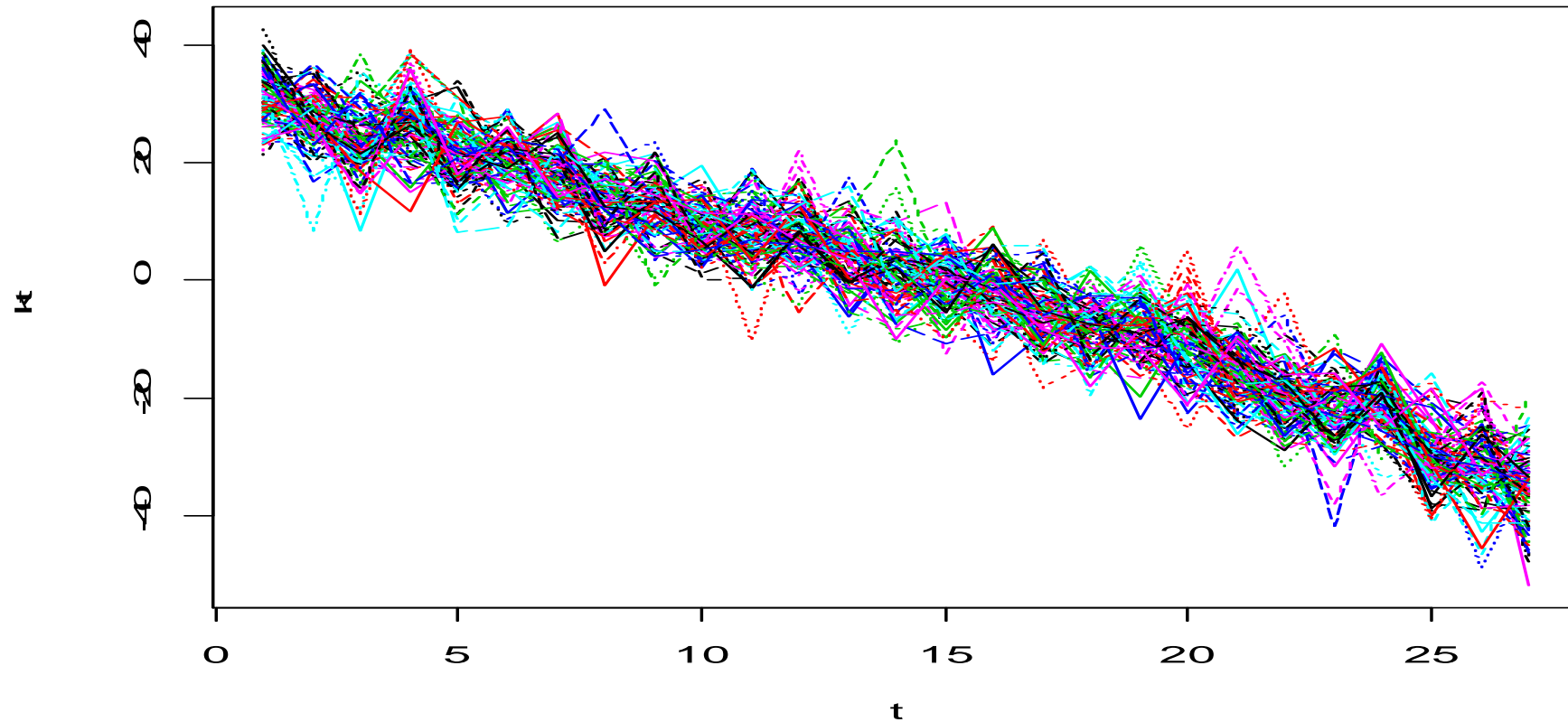


Figure 9 - Simulated paths for kt – Sieve Bootstrap

Numerical Application

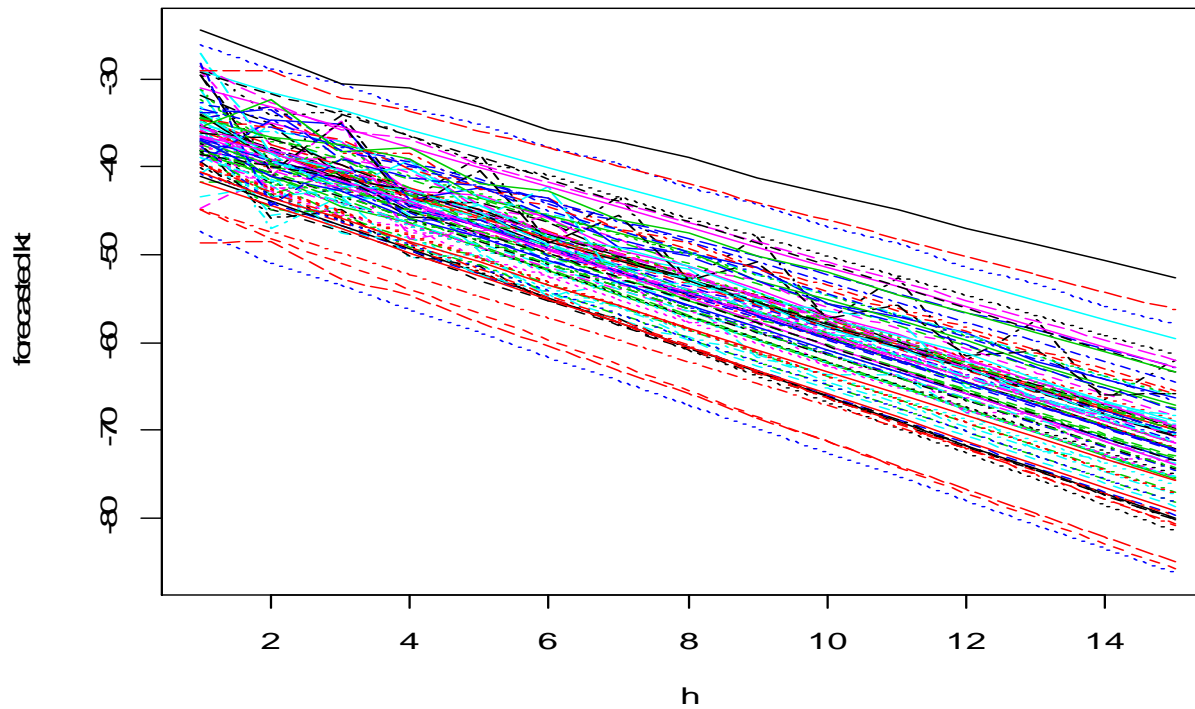


Figure 10 - Forecasted kt – Sieve Bootstrap

Numerical Application

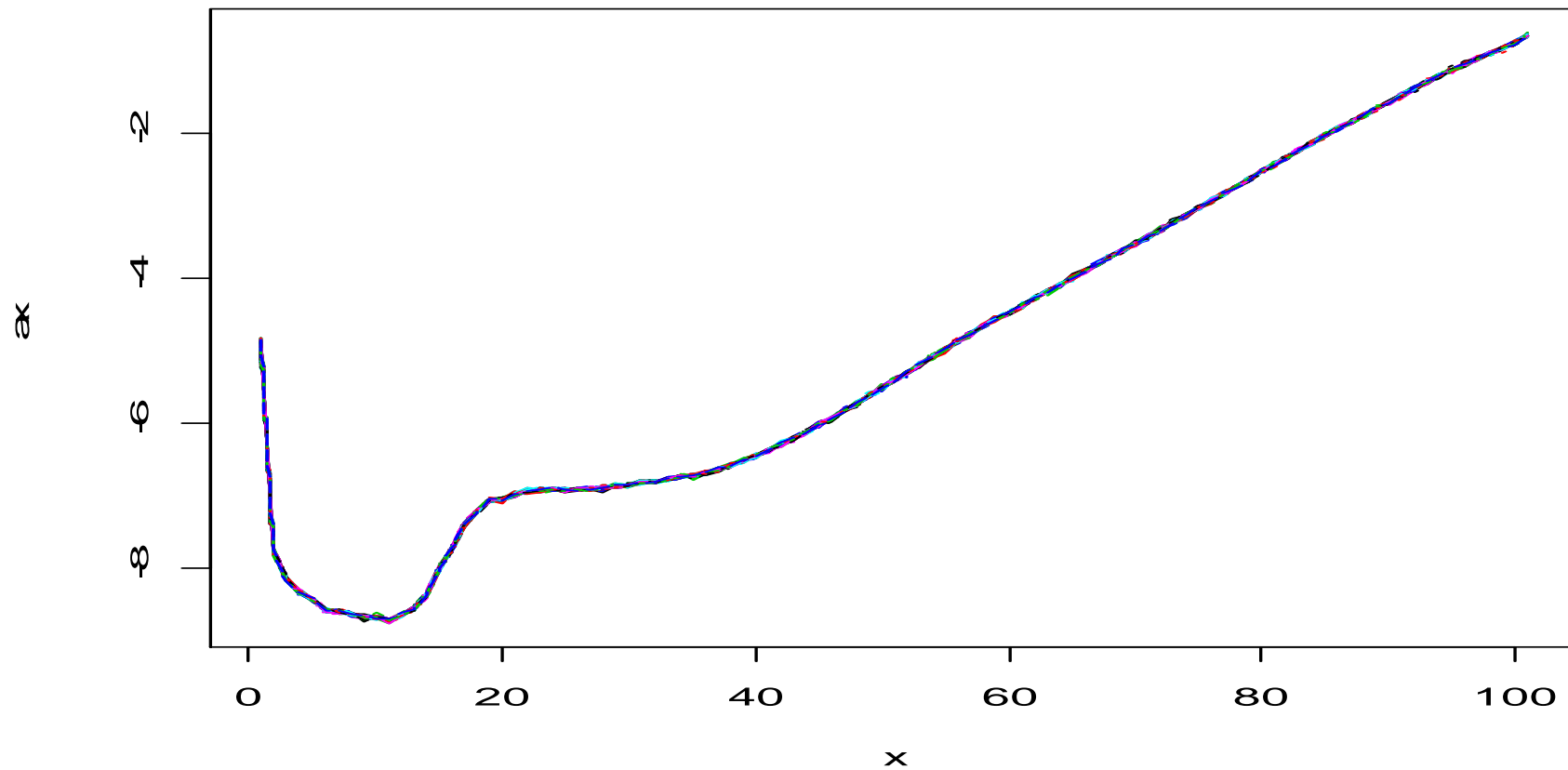


Figure 11 – Paths for ax – Residual standard bootstrap

Numerical Application

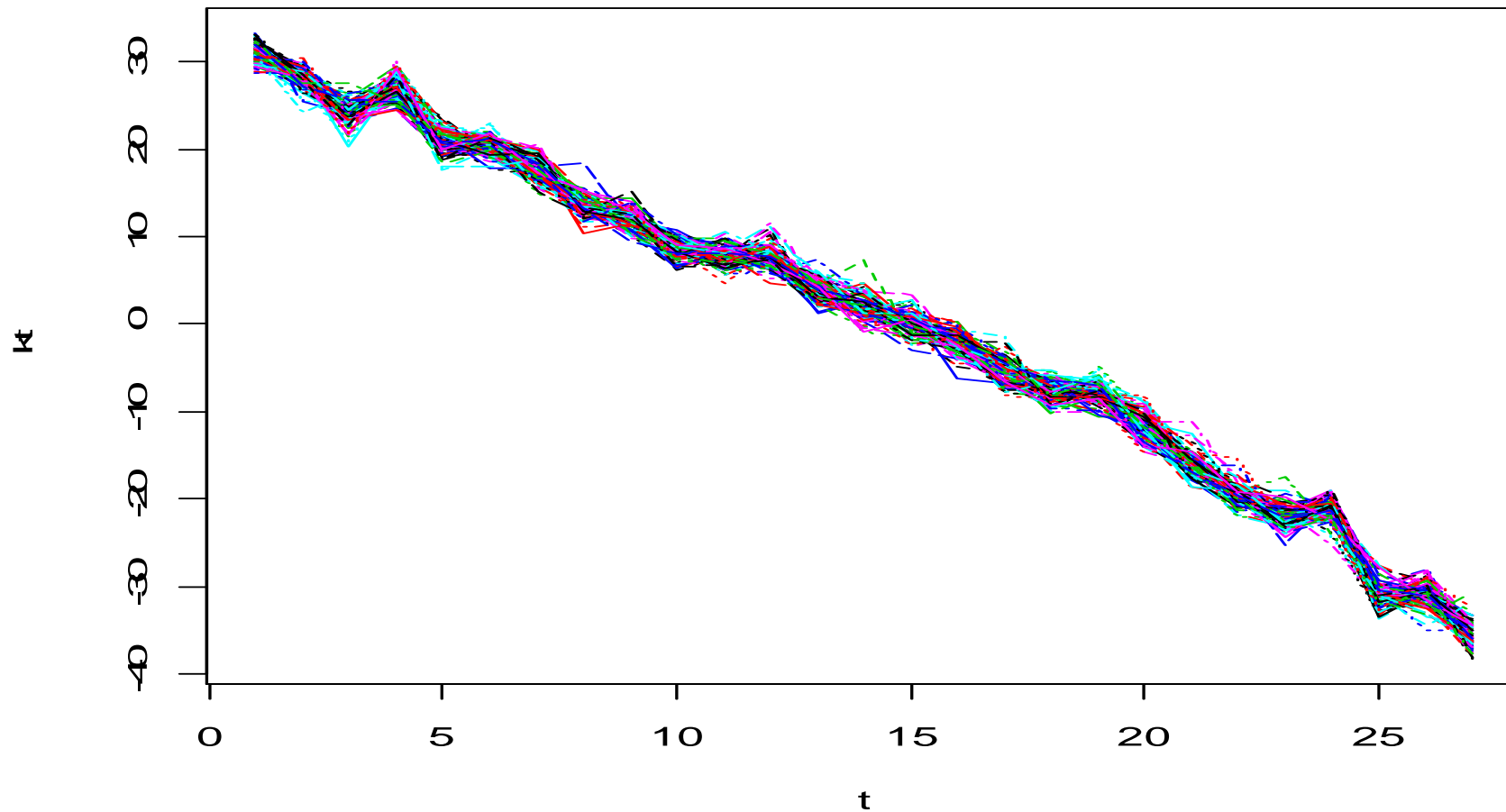


Figure 12 - Simulated paths for bx – Residual standard bootstrap

Numerical Application

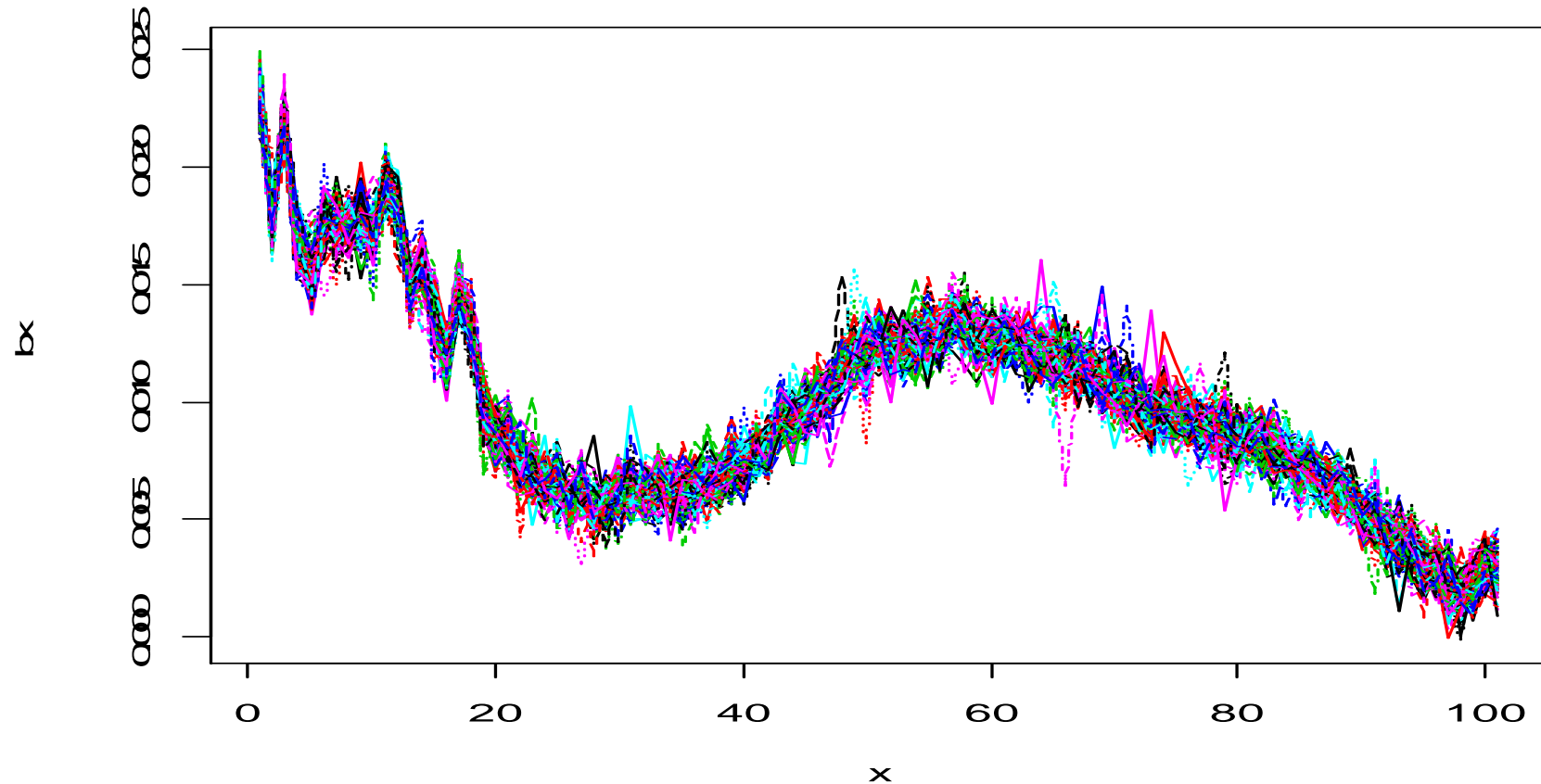


Figure 13 - Forecasted kt – Residual standard bootstrap

Numerical Application

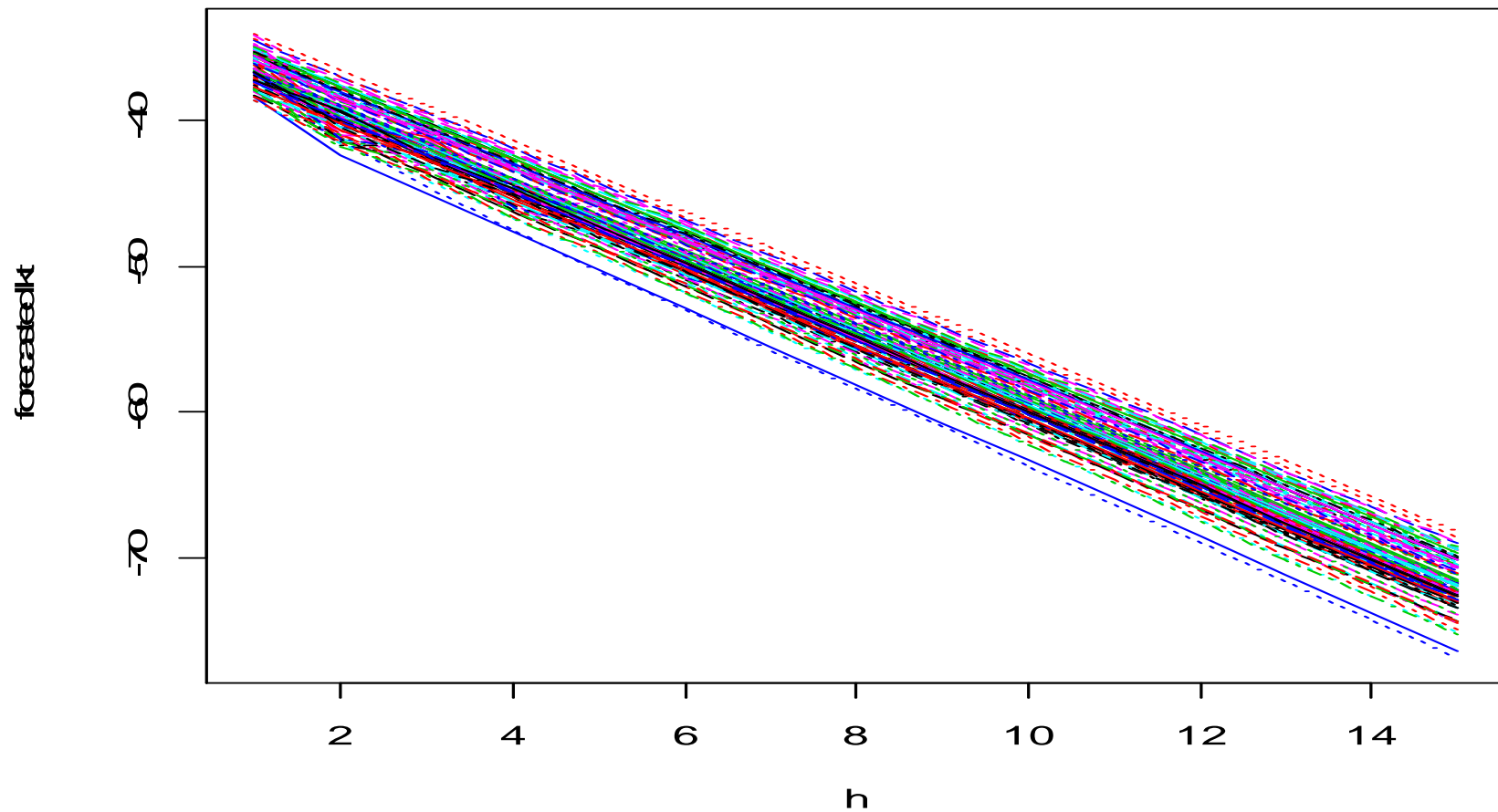


Figure 14 - Forecasted kt – Residual standard bootstrap

Numerical Application

h	Residual Bootstrap		Sieve Bootstrap	
	5%	95%	5%	95%
1	-38.10	-34.81	-41.75	-28.37
2	-41.45	-37.53	-45.66	-32.52
3	-43.66	-39.89	-47.82	-34.01
4	-46.30	-42.34	-49.82	-37.49
5	-48.82	-44.78	-52.95	-40.37
6	-51.40	-47.22	-55.07	-41.89
7	-53.95	-49.67	-57.69	-43.44
8	-56.52	-52.13	-61.43	-47.03
9	-59.08	-54.60	-63.52	-50.13
10	-61.66	-57.07	-66.63	-51.33
11	-59.55	-64.23	-68.97	-53.24
12	-62.02	-66.80	-72.17	-55.72
13	-69.37	-64.47	-74.42	-57.84
14	-71.94	-66.91	-77.69	-60.33
15	-74.50	-69.36	-80.27	-62.98

Table 3- Non parametric standard bootstrap and Sieve bootstrap 5% and 95% Confidence Intervals for k_{t+h}

Concluding Remarks

Our research proposes a particular bootstrap methodology, the LC Sieve Bootstrap, for capturing the *dependence* in deriving *prediction intervals*, thus avoiding a systematic over-estimation or under-estimation of the amount of uncertainty in the parameter estimates, respectively if negative or positive dependence dominates.

Concluding Remarks

- The *standard residual bootstrap* procedure does not preserve the correlation structure in the data.
- The *sieve bootstrap*, on the other hand, captures the dependency structure, leading to more reliable uncertainty measurement

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