

Pricing extreme mortality risk in the wake of the COVID-19 pandemic

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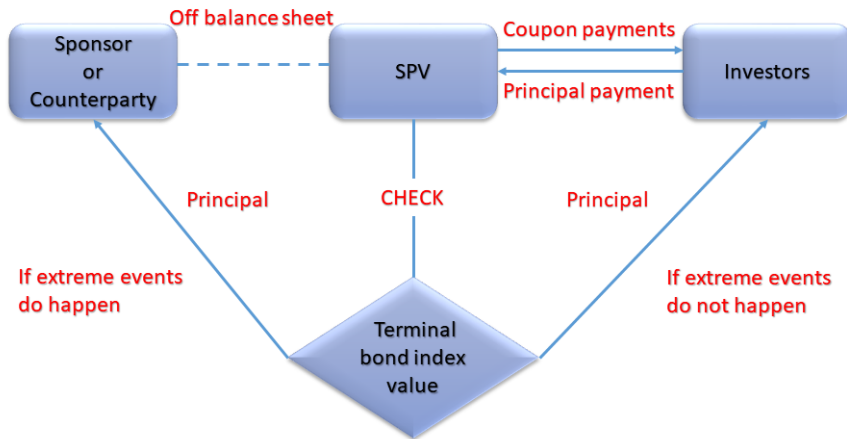
^[d] Smeal College of Business, **The Pennsylvania State University**

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Bayes Business School

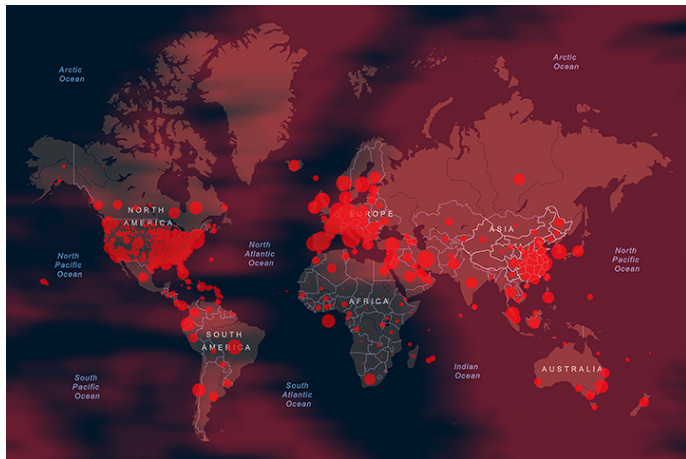
7–8 September 2023



Catastrophe bond and extreme insurance risk

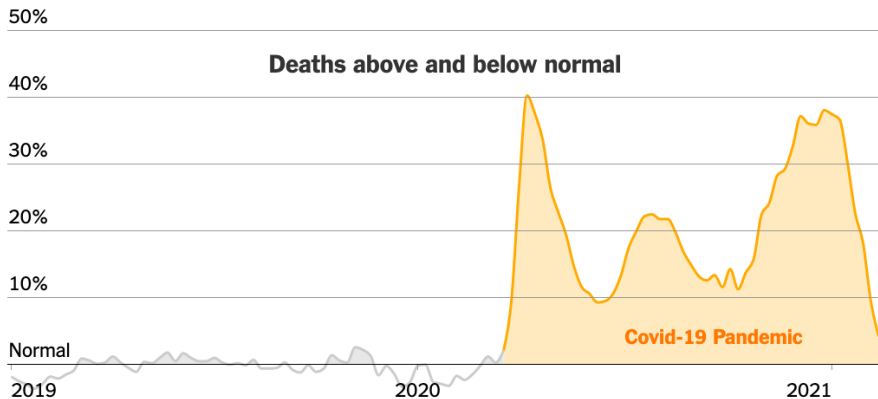


A 1 in 100 (?) year event



Source: <https://coronavirus.jhu.edu/map.html>

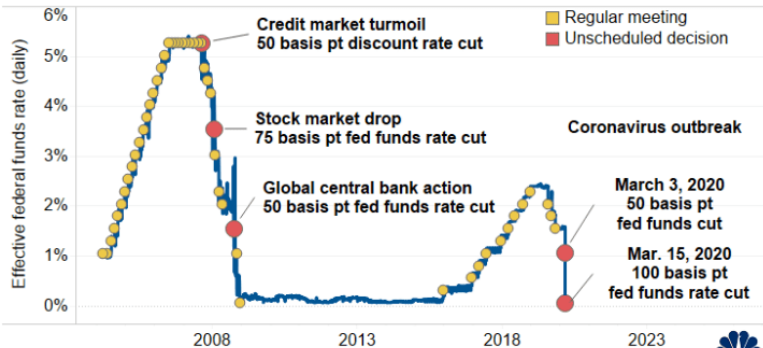
US: 574,000 more deaths than normal since March 2020



Source: <https://www.nytimes.com/interactive/2021/01/14/us/covid-19-death-toll.html>

Federal Reserve: emergency rate cut

Fed rate moves



SOURCE: Federal Reserve, New York Fed, St. Louis Fed

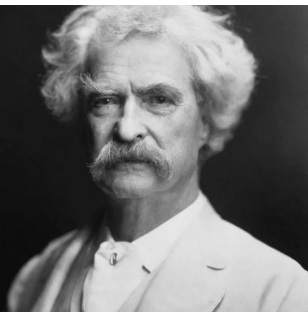


It is a common practice, but ...

It is **commonly** assumed, but **rarely** tested, that the interest rate and the mortality rate are *mutually independent*.

It ain't what you don't know
that gets you into trouble. It's
what you know for sure that
just ain't so.

Mark Twain



Outline of the talk

- ① Background and motivation
- ② The affine jump-diffusion (AJD) model
 - ▶ Dynamics under P
 - ▶ The pricing measure Q
- ③ Modeling mortality-linked securities
 - ▶ Model calibration
 - ▶ Atlas IX Capital Ltd. bond 2013
- ④ Implied market prices of risk
 - ▶ Numerical analysis
 - ▶ Sensitivity test
- ⑤ Conclusion remarks



The affine jump-diffusion (AJD) model

Assume that the bivariate process $\{\mathbf{Y}_t = (r_t, \mu_t)^\top\}_{0 \leq t \leq T}$ follows an affine jump-diffusion (AJD) process. Precisely, for $0 \leq t \leq T$,

$$d\mathbf{Y}_t = \mathbf{K}_t(\boldsymbol{\theta}_t - \mathbf{Y}_t)dt + \boldsymbol{\Sigma}_t\sqrt{\mathbf{S}_t}d\mathbf{W}_t + \sum_{i=1}^m d\mathbf{J}_{i,t}, \quad (1)$$

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- $\mathbf{J}_{i,t} = \sum_{k=1}^{N_{i,t}} \mathbf{X}_{i,k}$ is a **compound Poisson process** with rate $\lambda_i > 0$ and jump size distribution G_i .



The pricing measure Q

Under the measure Q , we look at two types of risk embedded in our model

- **Systematic risk:** $\{\mathbf{W}_t\}$, the main driving force inherent in the market

$$d\mathbf{W}_t^Q = d\mathbf{W}_t - \mathbf{\Gamma}_t dt, \quad 0 \leq t \leq T,$$

- **Jump risk:** For each $i = 1, \dots, m$, $\{\mathbf{J}_{i,t}\}_{0 \leq t \leq T}$ is a compound Poisson process with intensity λ_i^* and common jump size distribution G_i^* ;

$\{\mathbf{W}_t^Q\}_{0 \leq t \leq T}$ and $\{\mathbf{J}_{i,t}\}_{0 \leq t \leq T}$, $i = 1, \dots, m$, are mutually independent.



An important special case

The interest rate and mortality rate intensities are modeled by

$$\begin{cases} dr_t = (m_1 - d_1 r_t)dt + \sigma_1 dW_{1,t} + d \sum_{i=1}^{N_t} X_{1,i}, \\ d\mu_t = (m_2 - d_2 \mu_t)dt + \sigma_2 \left(\rho_1 dW_{1,t} + \sqrt{1 - \rho_1^2} dW_{2,t} \right) + d \sum_{i=1}^{N_t} X_{2,i}, \end{cases}$$

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- $m_i \in \mathbb{R}$, $d_i \neq 0$, $\sigma_i > 0$, for $i = 1, 2$, and $\rho_1 \in [-1, 1]$ are constants;

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- $\{N_t\}_{0 \leq t \leq T}$ and $\{\mathbf{X}_j\}_{j \in \mathbb{N}}$ are independent.

An important special case

Under the Q measure, we consider

- **Market price of diffusion risk:** $\Gamma_t = (\gamma_1, \gamma_2)^\top \in \mathbb{R}^2$, where

$$dW_{i,t} = dW_{i,t}^Q + \gamma_i dt \text{ for } i = 1, 2$$

We can then rewrite the dynamics as

$$\begin{cases} dr_t = (m_1^* - d_1 r_t) dt + \sigma_1 dW_{1,t}^Q + d \sum_{i=1}^{N_t} X_{1,i}, \\ d\mu_t = (m_2^* - d_2 \mu_t) dt + \sigma_2 \left(\rho_1 dW_{1,t}^Q + \sqrt{1 - \rho_1^2} dW_{2,t}^Q \right) + d \sum_{i=1}^{N_t} X_{2,i}, \end{cases}$$

with

$$m_1^* = m_1 + \gamma_1 \sigma_1, \quad m_2^* = m_2 + \gamma_1 \sigma_2 \rho_1 + \gamma_2 \sigma_2 \sqrt{1 - \rho_1^2}.$$



An important special case

Under the Q measure, we also consider

- **Market price of jump-frequency risk**

- ▶ $\{N_t\}_{0 \leq t \leq T}$ is a Poisson process with intensity $\lambda^* > 0$;
- ▶ $\chi = \frac{\lambda^*}{\lambda} > 0$ reflects the market price of jump-frequency risk.

- **Market price of jump-size risk:**

- ▶ Normalized multivariate exponential tilting to construct the common distribution G^* of $\{\mathbf{X}_j\}_{j \in \mathbb{N}}$;
- ▶ $\mathbf{X} \sim N(\nu_1^*, \nu_2^*; \phi_1, \phi_2; \rho_2)$ where $\nu_1^* = \nu_1 + \phi_1 \kappa_1$ and $\nu_2^* = \nu_2 + \phi_2 \kappa_2$.



Data description

We consider the US weekly mortality and interest rate data for the period of 2017–2020, which are collected from three sources as follows:

- **CDC COVID-19 Data.** We collect national-level weekly observed deaths and the expected deaths for the period Jan 2017–Dec 2020.
- **U.S. Census Bureau.** We collect population data for 2017–2020.

We define excess mortality as

$$\mu_t = \frac{d_t - E(d_t)}{e_t},$$

where d_t is the **observed number of deaths**, and $E(d_t)$ and e_t are, respectively, the **expected number of deaths** and the **population exposure**.



Data description

- ***Federal Reserve Economic Data (FRED)***. The weekly interest rate data comes from the 3-month treasury bill rates, collected at the same frequency and for the same period as the mortality data.

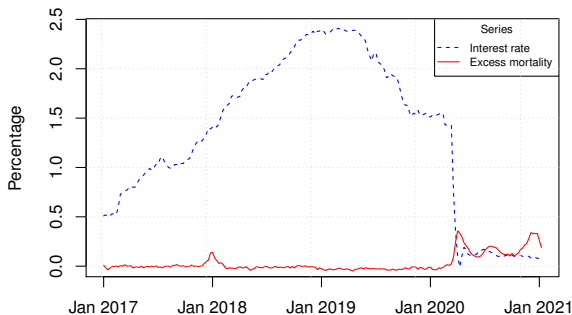


Figure: U.S. weekly interest rate and excess mortality

Model calibration: MCMC

- Likelihood based inference using MCMC.
- Evolution of the moment generating function is a Partial Differential Equation (PDE).
- This PDE is used to approximate the likelihood to a high degree of accuracy.
- Random Walk Metropolis Hastings used to explore the posterior.



Calibration including the pandemic experience

The interest rate and mortality rate intensities are modeled by

$$\begin{cases} dr_t = (0.005 - 0.126r_t)dt + 0.002dW_{1,t} + d \sum_{i=1}^{N_t} X_{1,i}, \\ d\mu_t = (0.002 - 2.301\mu_t)dt + 0.124(-0.038dW_{1,t} + 0.99928dW_{2,t}) + d \sum_{i=1}^{N_t} X_{2,i}, \end{cases}$$

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Calibration excluding the pandemic experience

The interest rate and mortality rate intensities are modeled by

$$\begin{cases} dr_t = (0.016 - 0.727r_t)dt + 0.002dW_{1,t} + d \sum_{i=1}^{N_t} X_{1,i}, \\ d\mu_t = (-0.217 - 16.368\mu_t)dt + 0.095 (0.017dW_{1,t} + 0.99986dW_{2,t}) + d \sum_{i=1}^{N_t} X_{2,i}, \end{cases}$$

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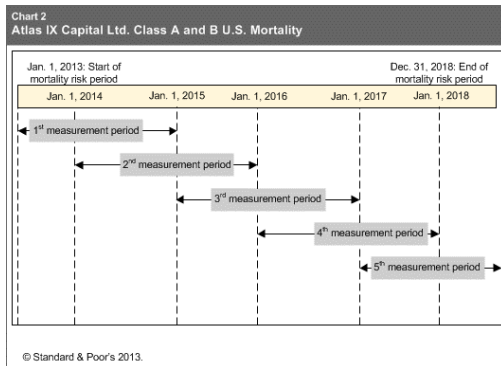
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Modeling mortality-linked securities

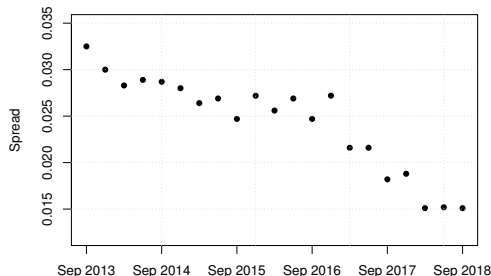
Launched by Atlas IX Capital Ltd in September 2013, the Atlas bond is the first catastrophe bond from SCOR which covers extreme mortality risk.



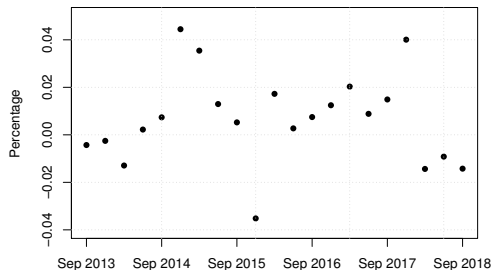
- **Risk period:** Jan 2013 – Dec 2018
- **Coupon payment:** 3.25% above the three-month LIBOR rate
- **Underlying mortality:** U.S. total population

Market prices of risk

Baseline price: the Atlas bond market price throughout 2013–2018.



Quarterly market-indicated spreads of the Atlas bond, published by Lane Financial L.L.C.



Estimated quarterly excess mortality.

Three scenarios

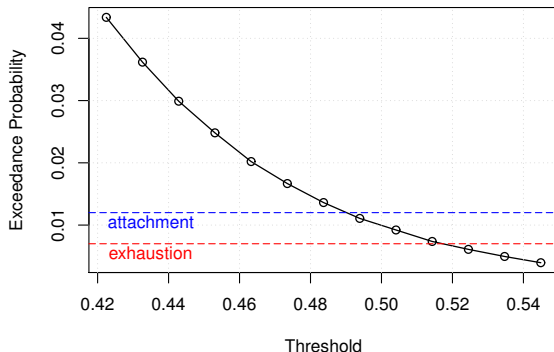
Idea: Minimizing the difference between theoretical price and observed price.

Method: Limited-memory Broyden–Fletcher–Goldfarb–Shanno method.

We compute the MPRs for three different scenarios:

- S1.** Underlying risks \Rightarrow post-pandemic model
Bond trigger levels \Rightarrow post-pandemic model;
- S2.** Underlying risks \Rightarrow pre-pandemic model;
Bond trigger levels \Rightarrow pre-pandemic model;
- S3.** Underlying risks \Rightarrow post-pandemic model
Bond trigger levels \Rightarrow pre-pandemic model.

Scenario one

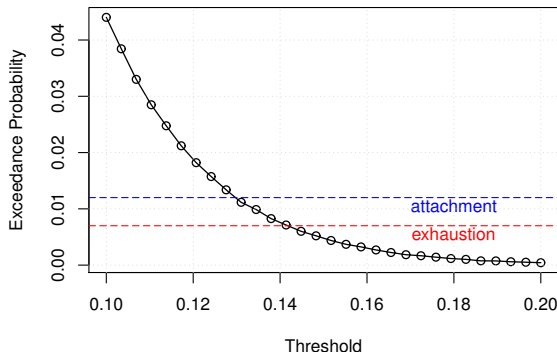


The MPR vector is obtained as

$$\zeta_1 = (\gamma_1, \gamma_2; \kappa_1, \kappa_2; \chi) = (0.4161, 0.1897; 0.1119, 0.3889; 1.0918).$$

Investors' perceived parameter values under Q are all riskier than their P measure counterparts.

Scenario two

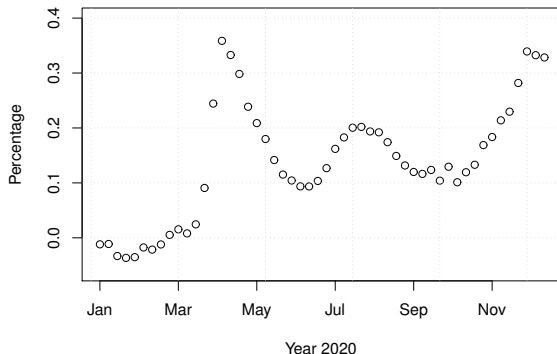


The MPR vector is obtained as

$$\zeta_2 = (\gamma_1, \gamma_2; \kappa_1, \kappa_2; \chi) = (0.1099, 0.0924; 0.1068, 0.0961; 1.3144).$$

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Scenario three



The MPR vector is obtained as

$$\zeta_3 = (\gamma_1, \gamma_2; \kappa_1, \kappa_2; \chi) = (1.0944, -1.6196; 1.0743, -4.0847; 1.6419).$$

Investors are receiving negative mortality risk premia although the interest rate risk premia they receive are positive.

Three scenarios

Parameter	Scenario 1		Scenario 2		Scenario 3	
	Under P	Under Q	Under P	Under Q	Under P	Under Q
$m_1 \leftrightarrow m_1^*$	0.005	0.0058	0.016	0.0162	0.005	0.0072
$m_2 \leftrightarrow m_2^*$	0.002	0.0235	-0.217	-0.2080	0.002	-0.2038
$\nu_1 \leftrightarrow \nu_1^*$	-0.001	-0.0008	0.000	0.0001	-0.001	0.0011
$\nu_2 \leftrightarrow \nu_2^*$	0.035	0.0638	0.026	0.0314	0.035	-0.2673
$\lambda \leftrightarrow \lambda^*$	4.865	5.3118	1.909	2.5091	4.865	7.9877

Sensitivity analysis

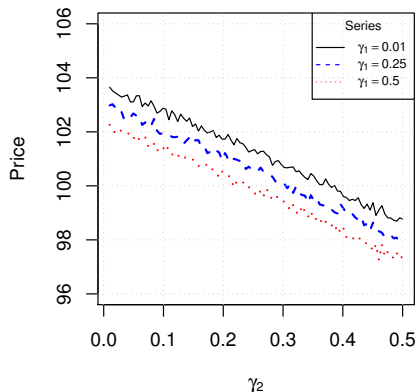
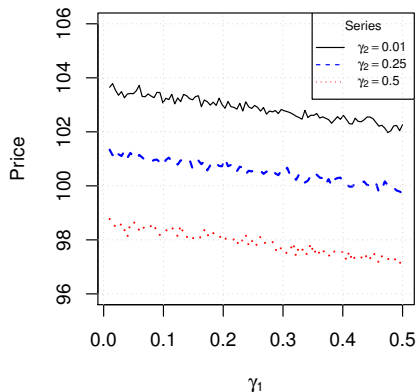


Figure: Sensitivity analysis with respect to γ_1 and γ_2

Sensitivity analysis

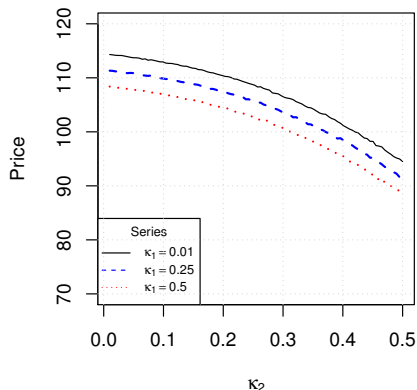
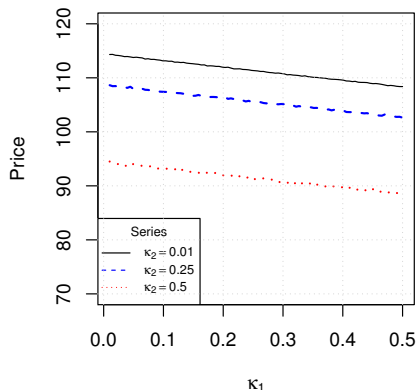


Figure: Sensitivity analysis with respect to κ_1 and κ_2

Sensitivity analysis

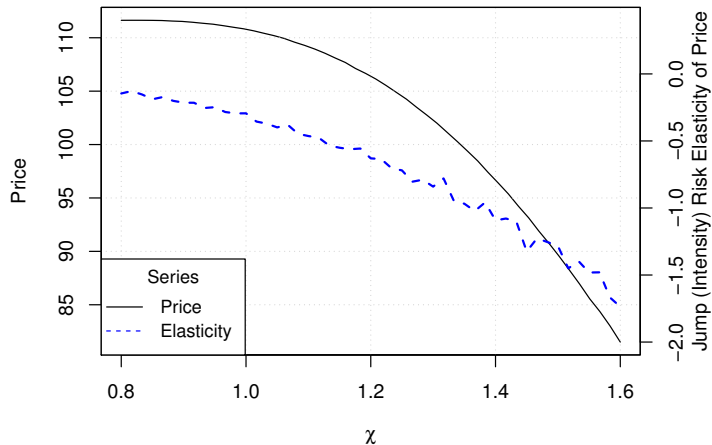


Figure: Sensitivity analysis with respect to χ

Concluding remarks

In this research, we ...

- **Propose** a bivariate AJD structure to jointly model the interest rate and excess mortality.
- **Show** that the COVID-19 pandemic experience greatly intensifies the negative instantaneous correlations.
- **Develop** a risk-neutral pricing measure that accounts for both a diffusion risk premium and a jump risk premium.
- **Solve** for the market prices of risk based on mortality CAT bond prices.



End of presentation

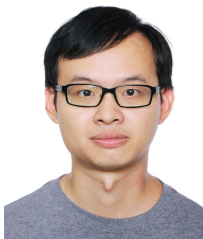
Thank you!

Any questions/ comments/ suggestions?

Contact email: han.li@unimelb.edu.au



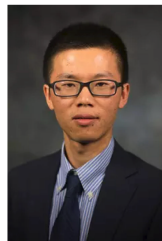
(a) Han Li



(b) Haibo Liu



(c) Qihe Tang



(d) Zhongyi Yuan



Appendix

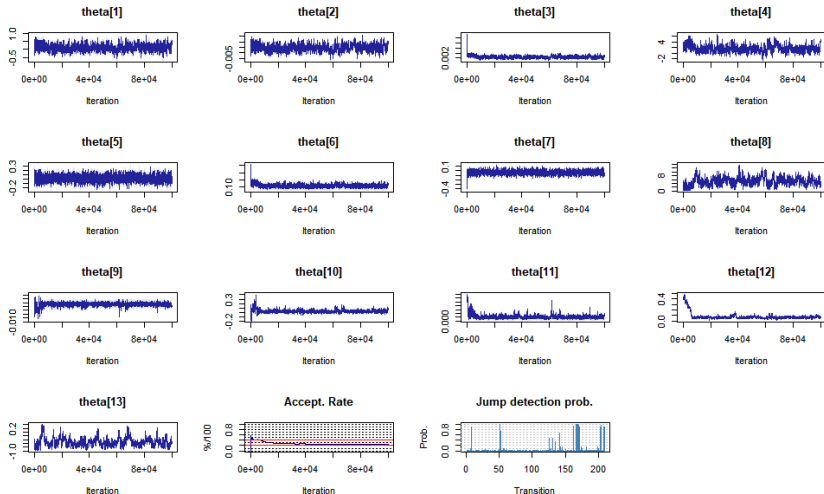


Figure: Convergence diagnostics for the MCMC sampler (including pandemic experience).

Appendix

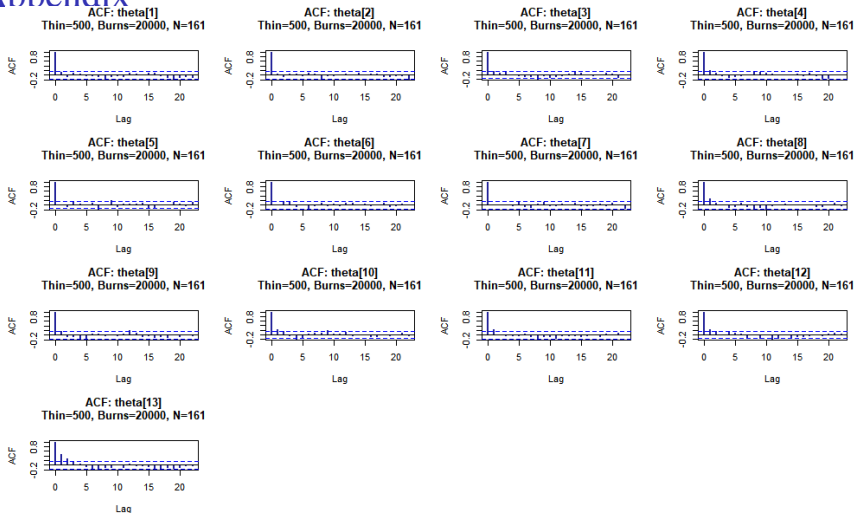


Figure: ACF of MCMC draws for each parameter in model (including pandemic experience).

Appendix

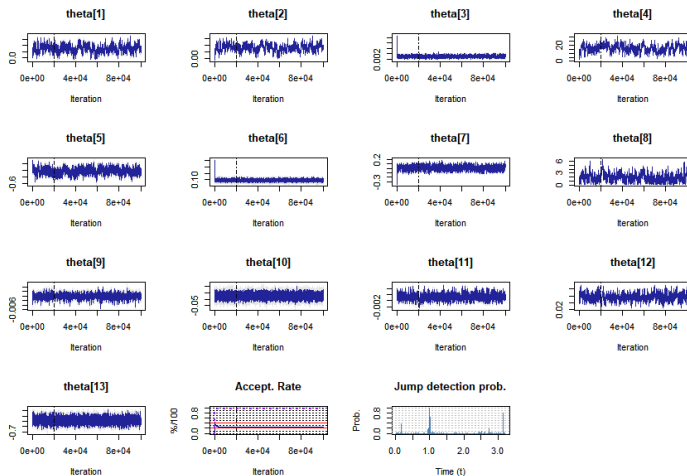


Figure: Convergence diagnostics for the MCMC sampler (excluding pandemic experience).

Appendix

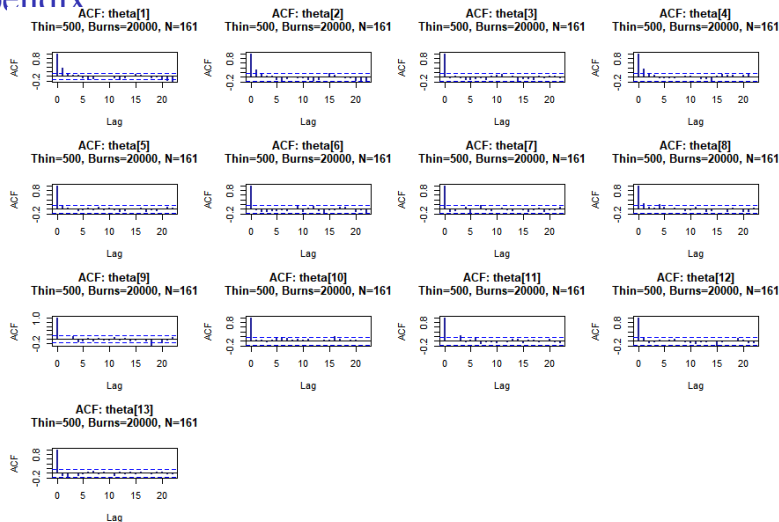


Figure: ACF of MCMC draws for each parameter in model (excluding pandemic experience).