



Max Planck Odense Center on the Biodemography of Aging
University of Southern Denmark

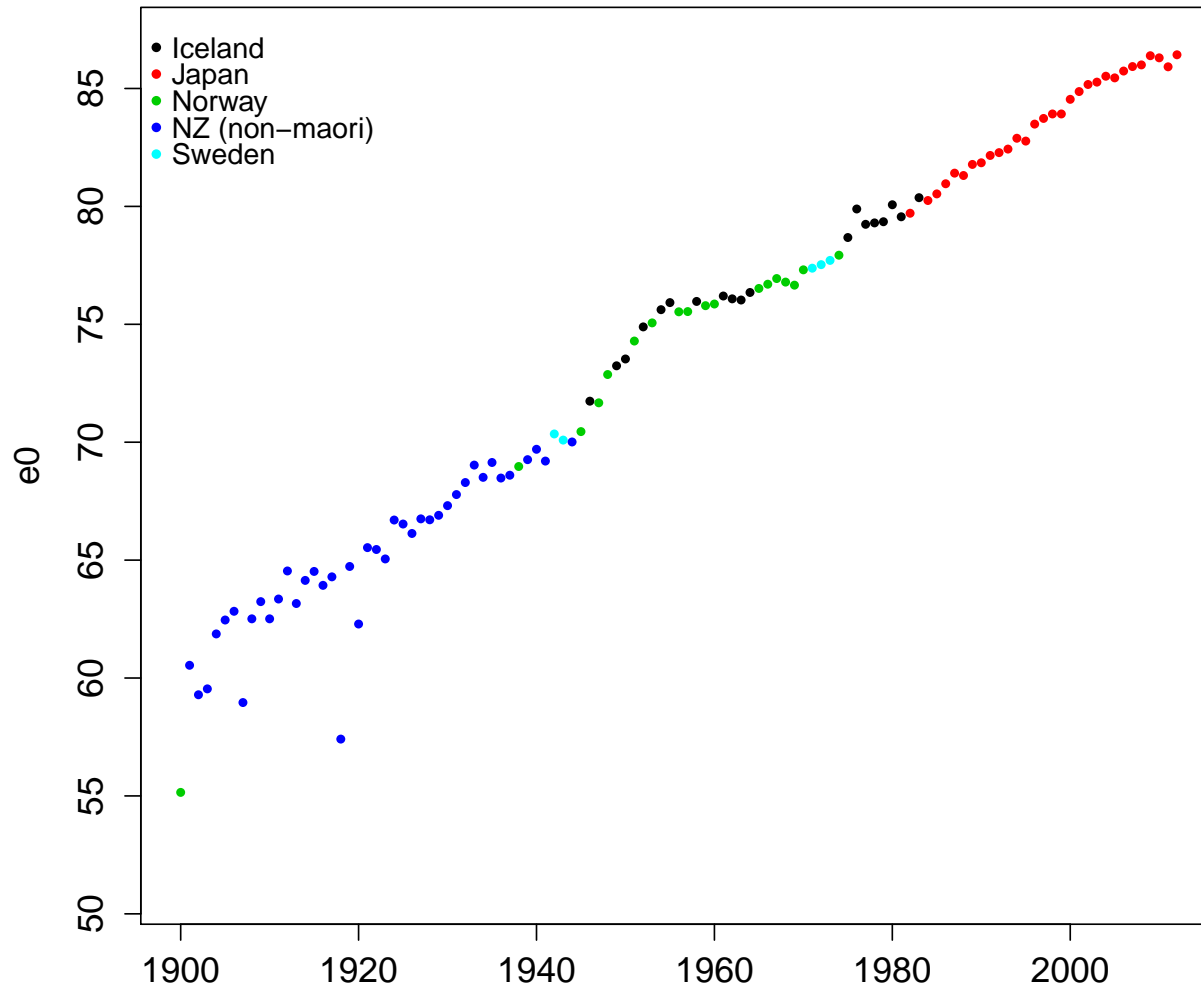
Projecting Maximum Country Life Expectancy using Provincial Data Only

Anthony Medford

Twelfth International Longevity Risk and Capital Markets
Solutions Conference

Overview

- ▶ What is Global Best practice Life expectancy?
- ▶ Extreme Value Theory in brief and its relation to Best practice Life expectancy
- ▶ Regional Best practice Life expectancy and inference

Females e_0 Female Best Practice e_0 

Global Best Practice Life Expectancy

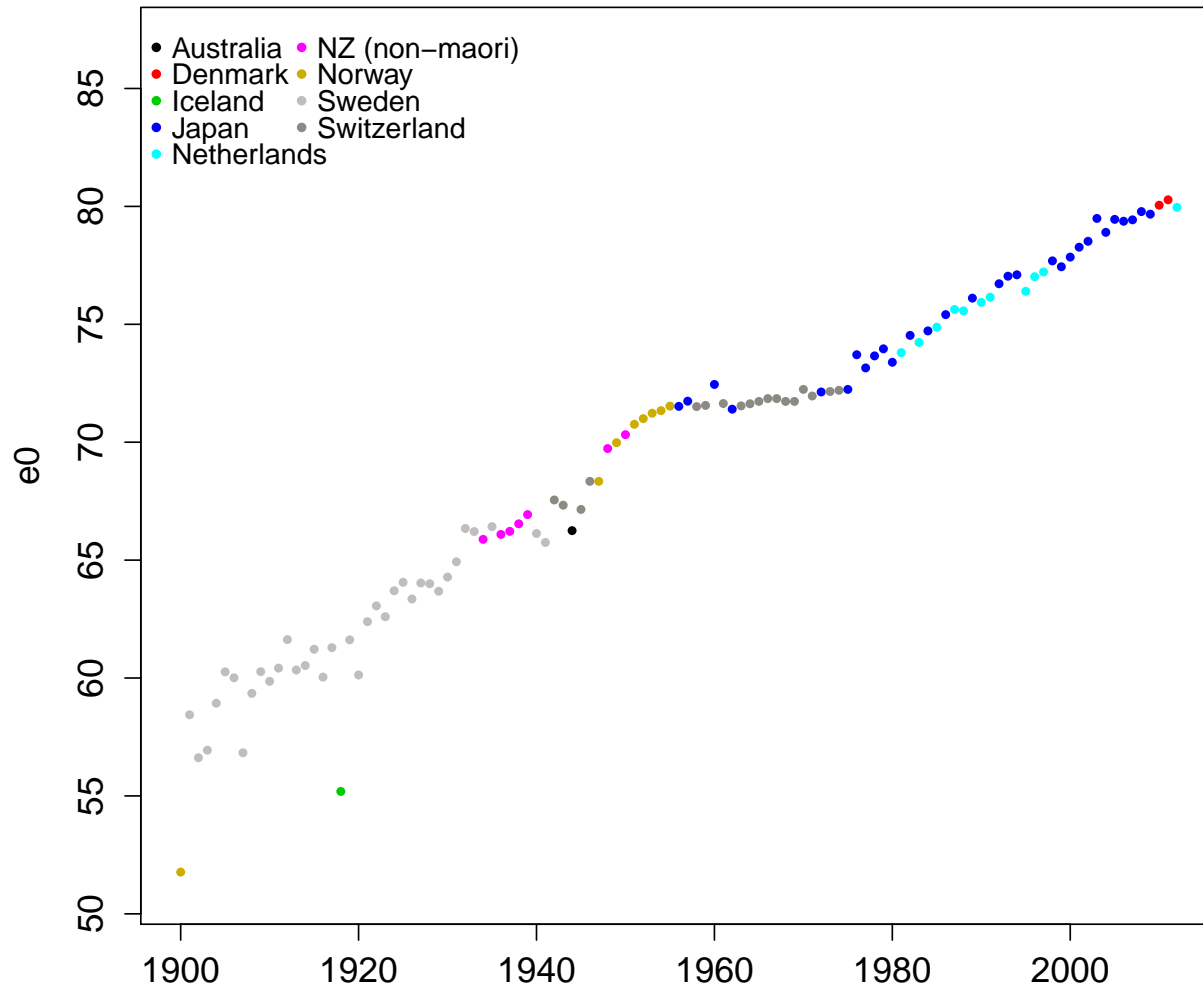
- ▶ Global Best Practice Life Expectancy (BPLE) is the maximum life expectancy observed among nations at a given age.
- ▶ At birth, has been increasing almost linearly - beginning in Scandinavia c. 1840 - at about 3 months per year (Oeppen and Vaupel, 2002).
- ▶ Life expectancy trends may fit better than individual-country trends in age-standardized (log) death rates (White, 2002).

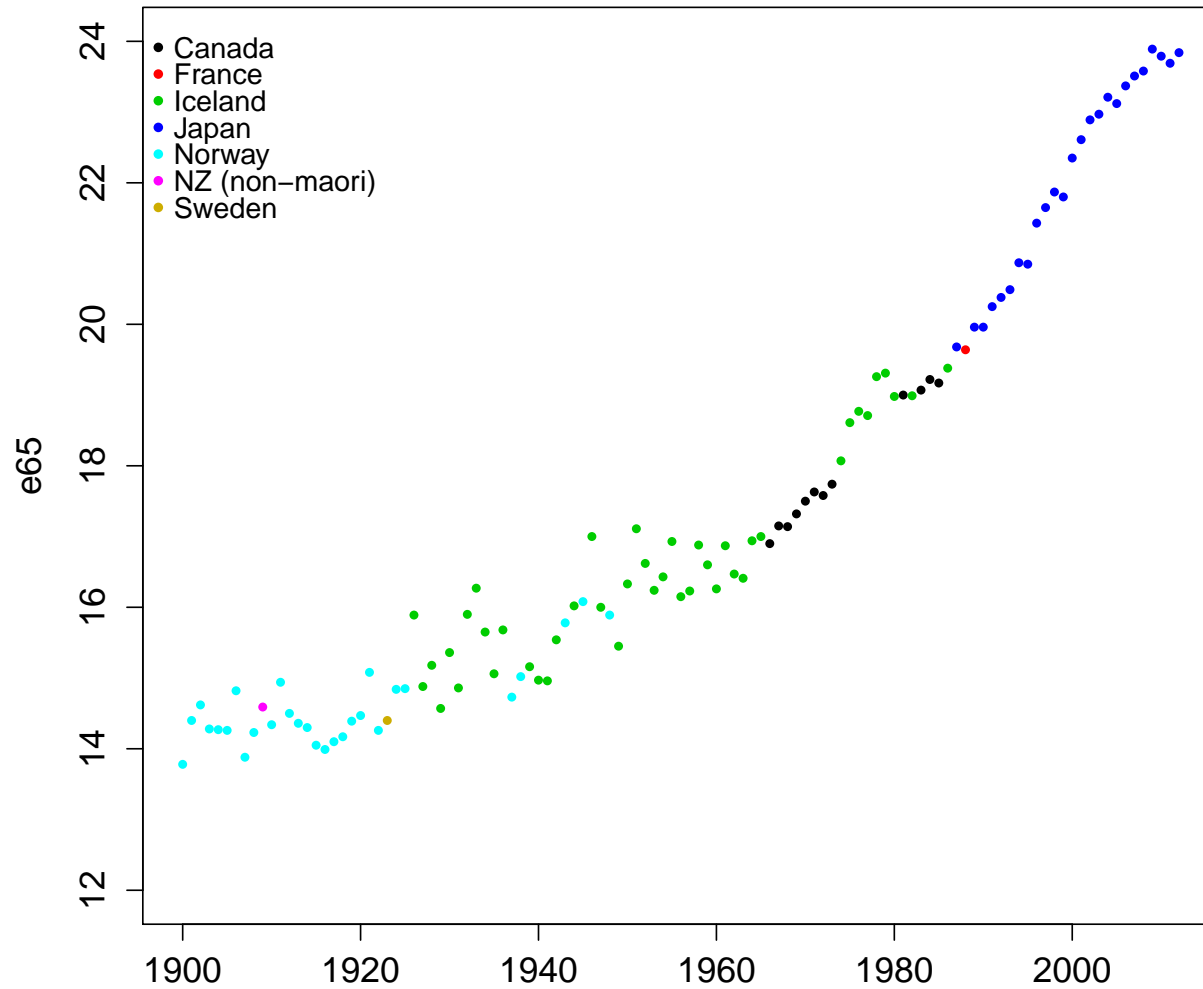
Global Best Practice Life Expectancy

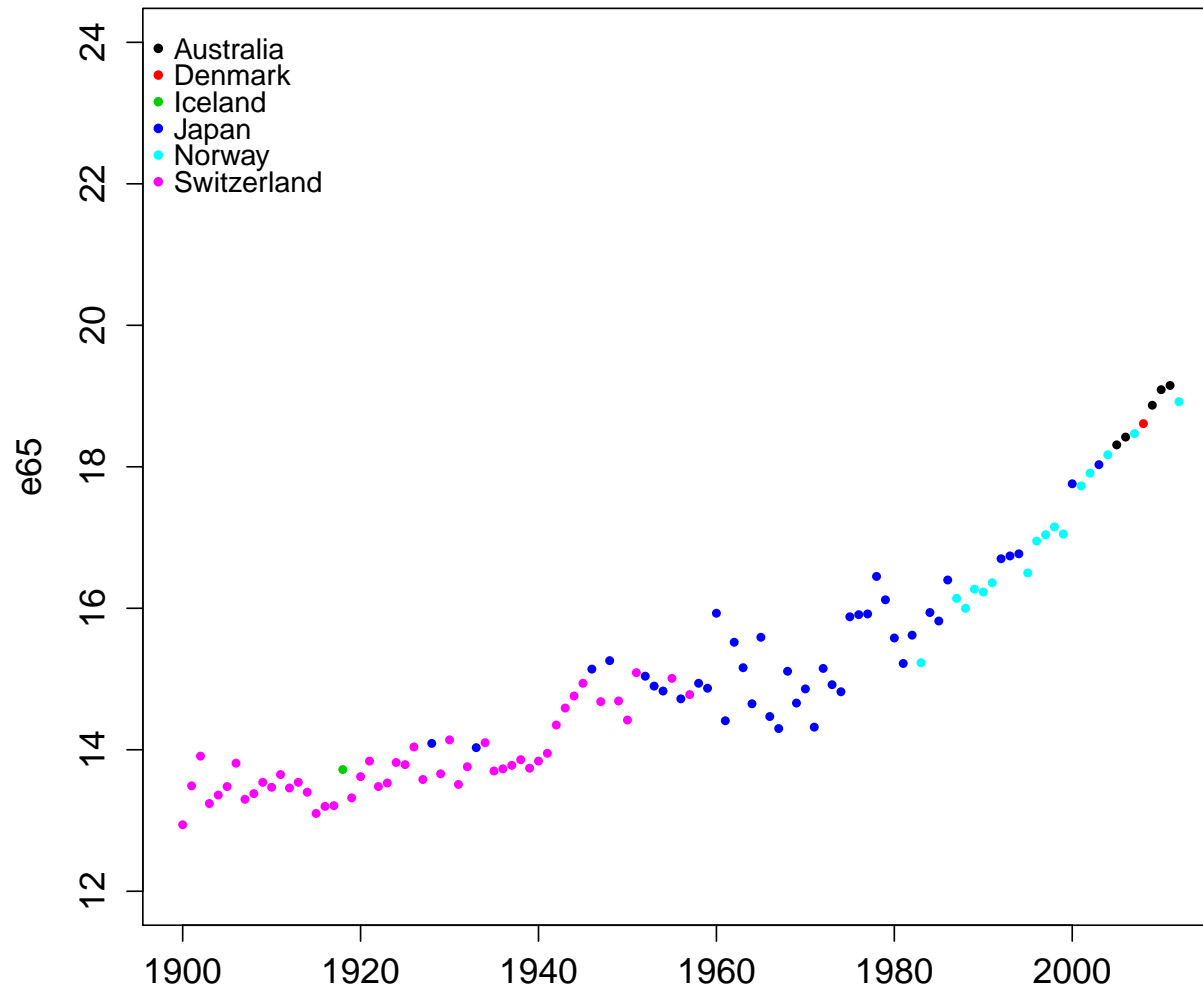
- ▶ Nations experience more rapid life expectancy gains when they are farther below BPLE and tend to converge towards BPLE (Torri and Vaupel, 2012).
- ▶ It is sensible to consider national mortality trends in a larger international context rather than individual projections (Lee, 2006; Wilmoth, 1998).

Males e_0

Male Best Practice e_0



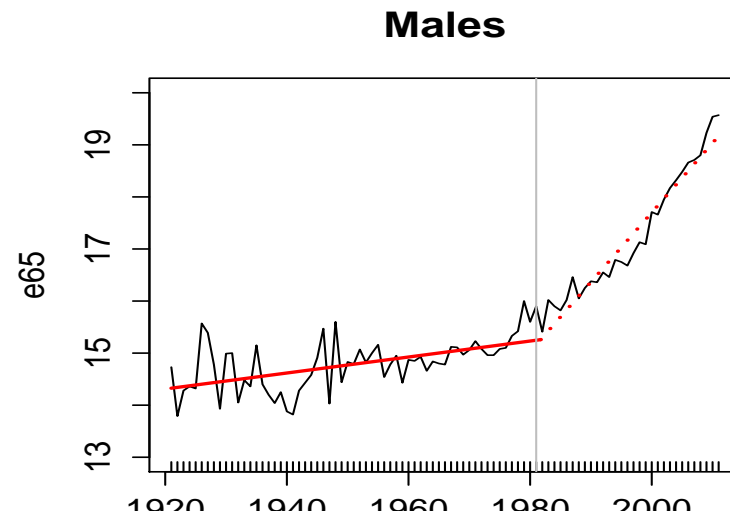
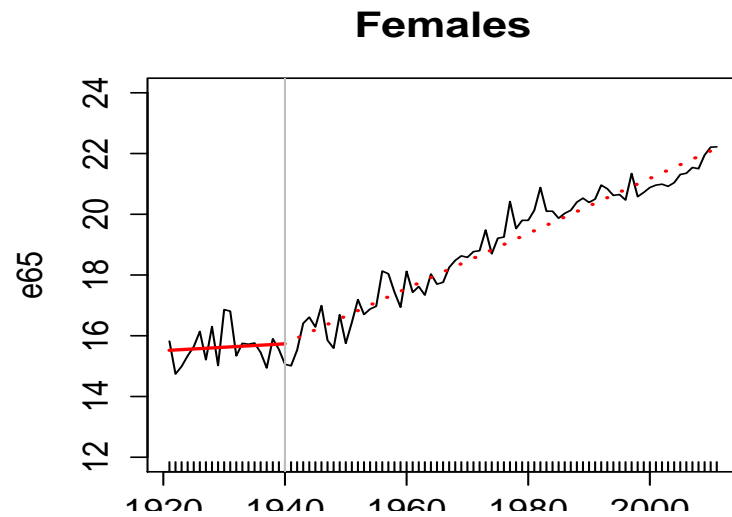
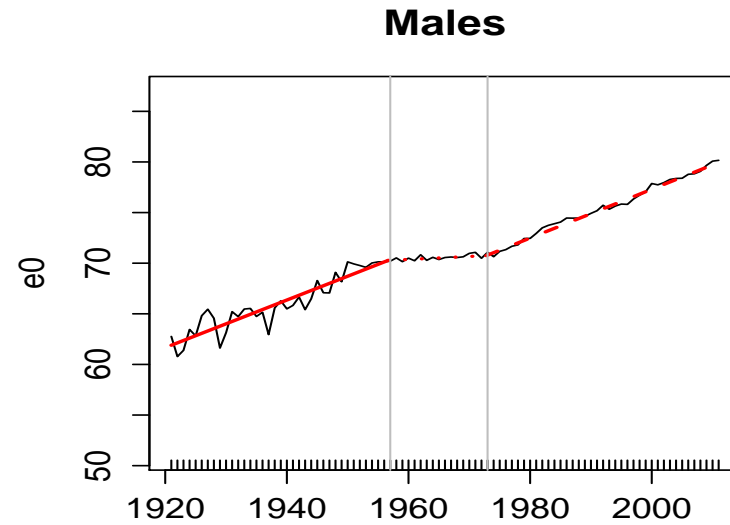
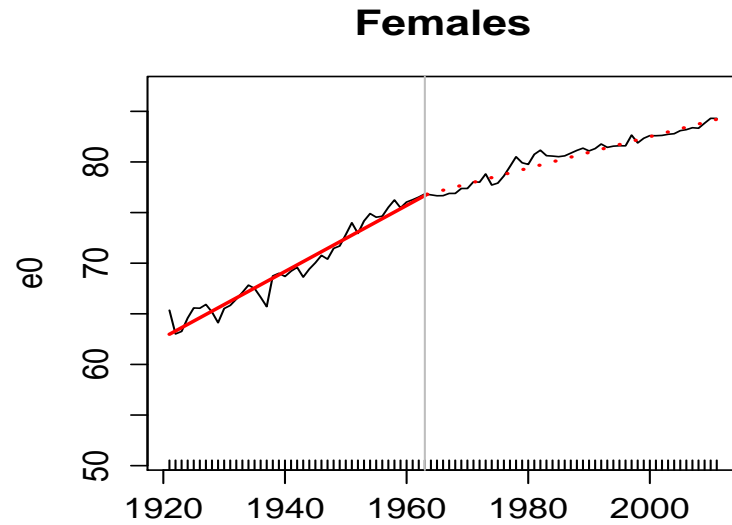
Females e_{65} Female Best Practice e_{65} 

Males e_{65} Male Best Practice e_{65} 

Breakpoints

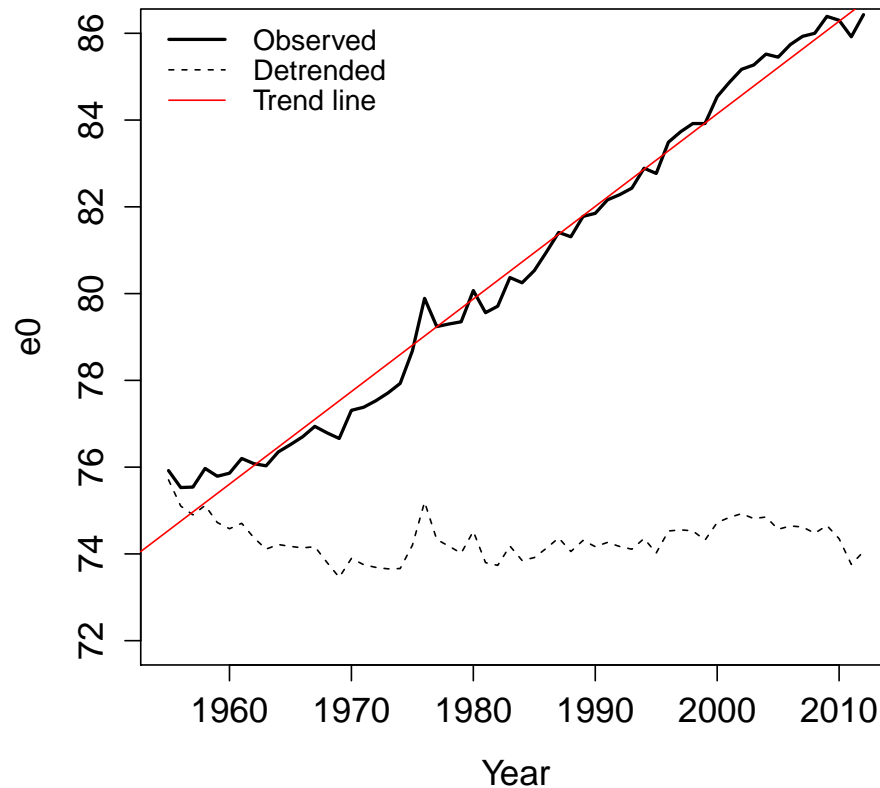
- ▶ Vallin and Meslé (2009) expanded on work of Oeppen and Vaupel and argued that BPLE trend may comprise multiple segments
- ▶ Each segment corresponds to distinct health transition phases

Breakpoints



Can we model Global BPLE?

Female Best Practice e0



Kernel Density and fitted GEV

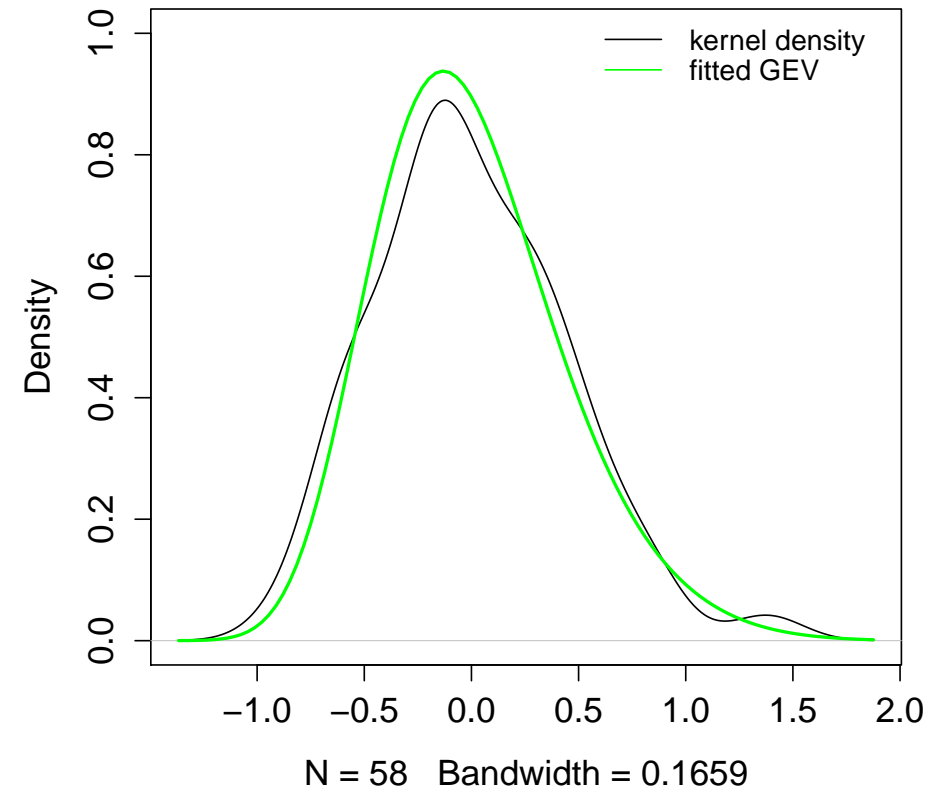


Figure: Left panel: raw and detrended data. Right panel: kernel density and fitted GEV distribution.

Can we model Global BPLE?

Suppose that X_1, X_2, \dots, X_n is a sequence of independent, identically distributed random variates all having a common distribution function $F(x)$.

Let $M_n = \max\{X_1, X_2, \dots, X_n\}$.

The distribution of the maxima, M_n , converges (for large n) to the Generalized Extreme Value (GEV) Distribution.

Theoretical motivation

Extremal Types theorem

If there exists sequences of constants $\{a_n > 0\}$ and $\{b_n\}$, such that as $n \rightarrow \infty$,

$$P\left(\frac{M_n - b_n}{a_n} \leq z\right) \rightarrow G(z) \quad (1)$$

where $G(z)$ is a non-degenerate distribution function, then G **must** be a member of the Generalized Extreme Value (GEV) family of distributions (Fisher and Tippett, 1928; Gnedenko, 1943).

Theoretical motivation

Extremal Types theorem

- ▶ This is a remarkable result because regardless of the underlying distribution, the distribution of the maxima (or minima) converges to one of the Generalized Extreme Value family of distributions.
- ▶ Can maximum period life expectancies be approximately modeled as a GEV?

The Generalized Extreme Value Distribution

$$G(z) = \exp\left\{-\left[1 + \xi\left(\frac{z - \mu}{\sigma}\right)\right]_+^{-\frac{1}{\xi}}\right\} \quad (2)$$

where $b_+ = \max(0, b)$. The situation where $\xi = 0$ is not defined in (2), but taken as the limit as $\xi \rightarrow 0$, given by

$$G(z) = \exp\left\{-\exp\left[-\left(\frac{z - \mu}{\sigma}\right)\right]\right\}. \quad (3)$$

The Generalized Extreme Value Distribution

- ▶ μ is the location parameter
- ▶ σ is the scale parameter
- ▶ ξ is the shape parameter, which determines the tail behaviour
 - $\xi > 0$: polynomial tail decay and the Fréchet Distribution
 - $\xi = 0$: exponential tail decay and the Gumbel Distribution
 - $\xi < 0$: bounded upper finite end point and the Weibull Distribution

Inference

Quantiles

Inverting the GEV distribution function:

$$z_p = \mu - \frac{\sigma}{\xi} \left[1 - \{-\log(1 - p)\}^{-\xi} \right],$$

where p is the tail probability and $G(z_p) = 1 - p$

Return Levels

- ▶ Simply a different way of thinking about the quantiles.
- ▶ If data are annual the $(1 - p)$ th quantile would be exceeded on average once every $1/p$ years.

Regional Best Practice Life Expectancy?

THE IDEA

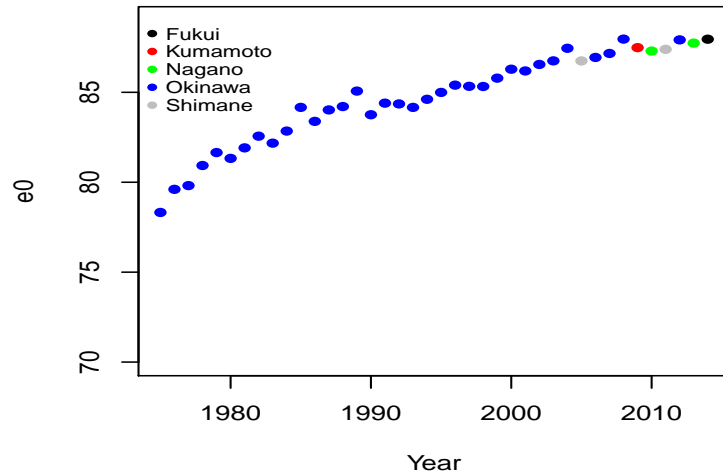
- ▶ Can the notion of BPLE be extended to regions smaller than the global whole?
- ▶ If we find BPLE over an arbitrary region - itself comprised of smaller subregions - would there also be a regular temporal evolution e.g. strong (piecewise) linear trends?
- ▶ What sort of inferences can we perform?

Data

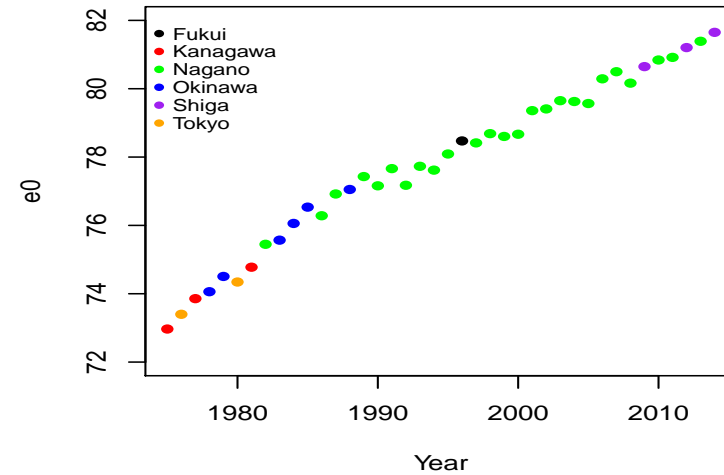
- ▶ Canadian Human Mortality Database (CHMD)
 - ▶ Life expectancy data broken down by province
 - ▶ Covers period from 1921 to 2011 (but Newfoundland from 1949)
- ▶ Japanese Mortality Database (JMD)
 - ▶ Life expectancy data broken down by prefecture
 - ▶ Covers period from 1975 to 2014

Japan - Maximum e_x by Prefecture

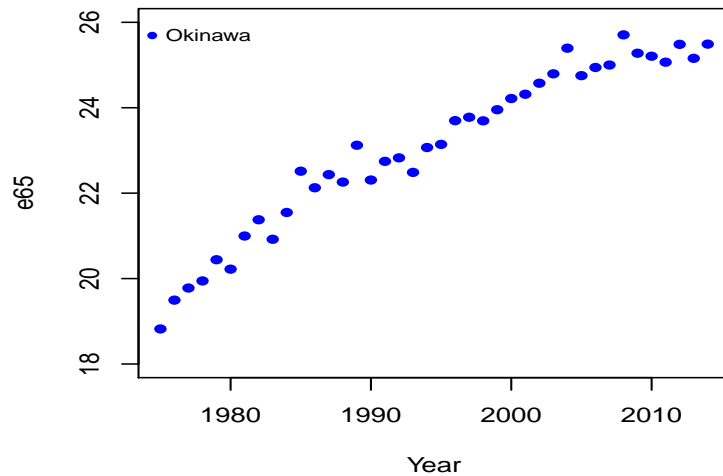
Female Best Practice e_0



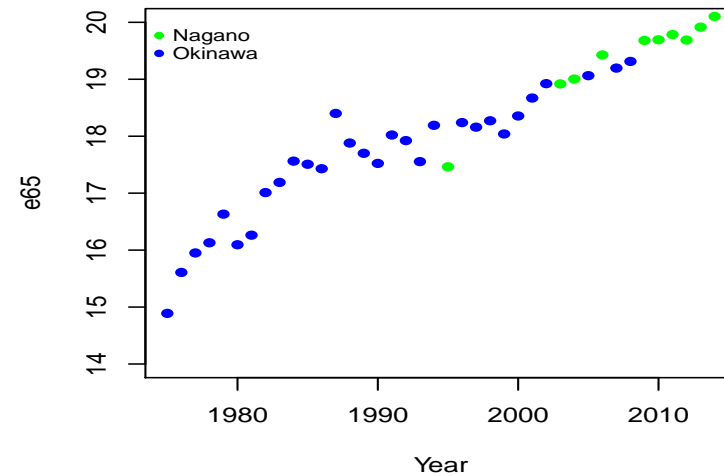
Male Best Practice e_0



Female Best Practice e_{65}

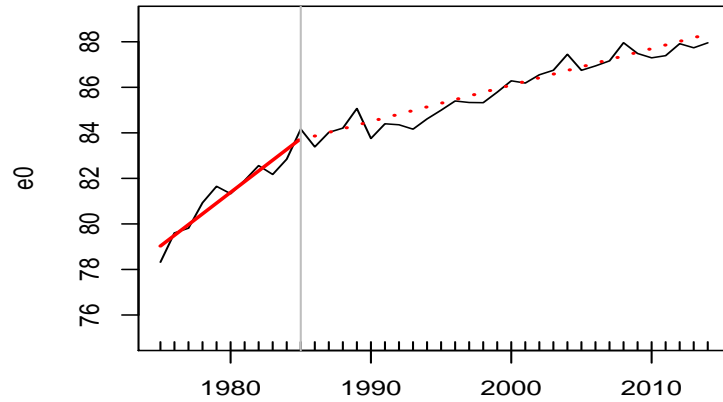


Male Best Practice e_{65}

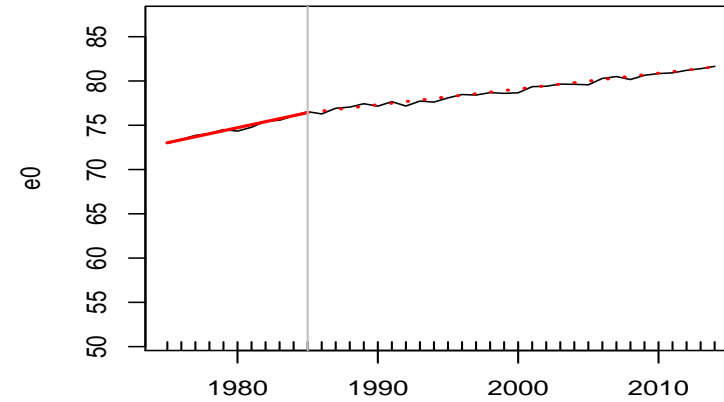


Japan - Maximum e_x by Prefecture

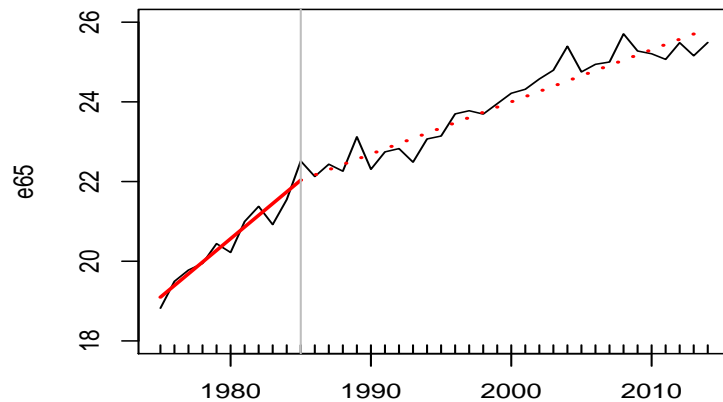
Females



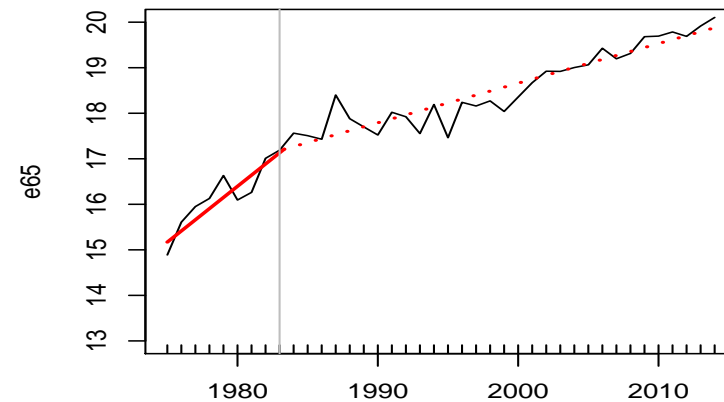
Males



Females

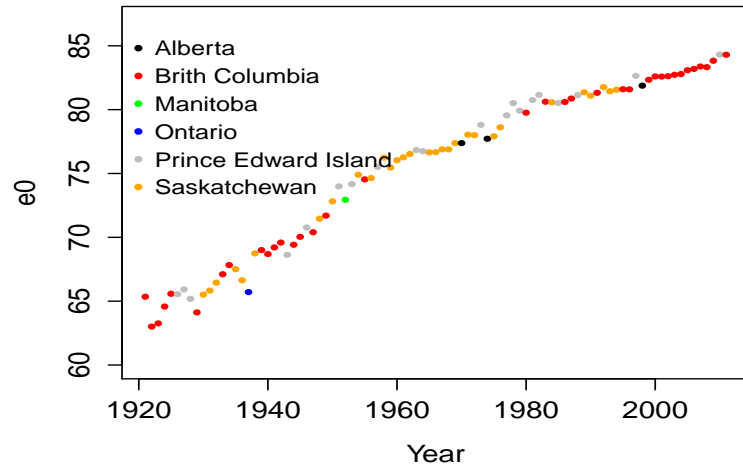


Males

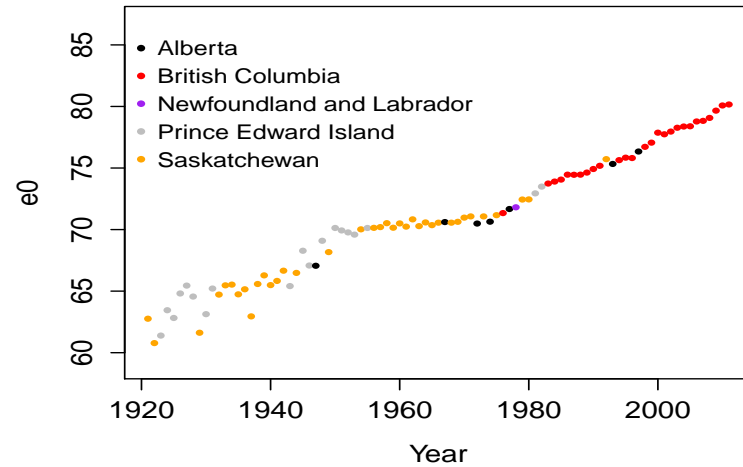


Canada - Maximum e_x by Province

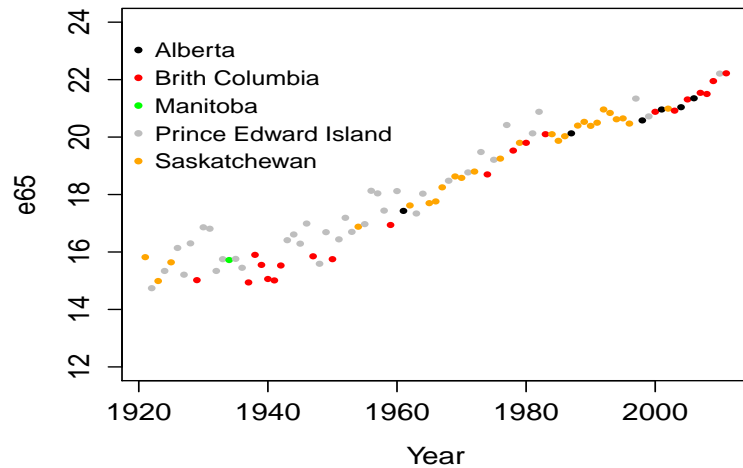
Female Best Practice e_0



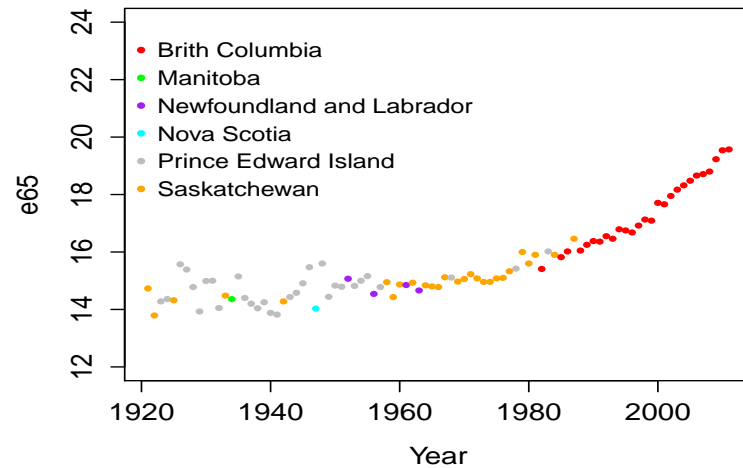
Male Best Practice e_0



Female Best Practice e_{65}

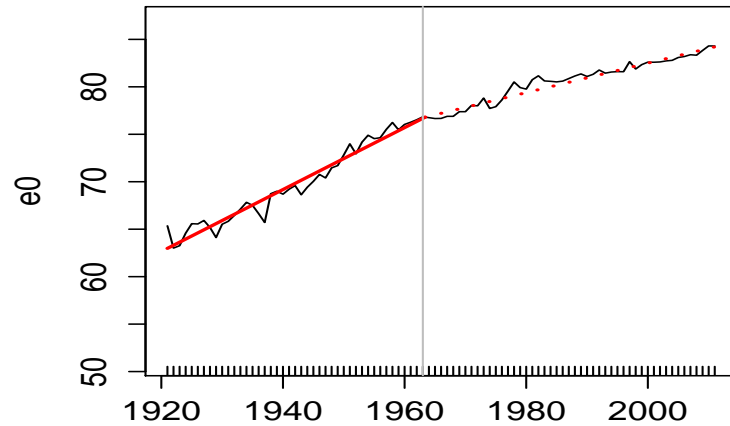


Male Best Practice e_{65}

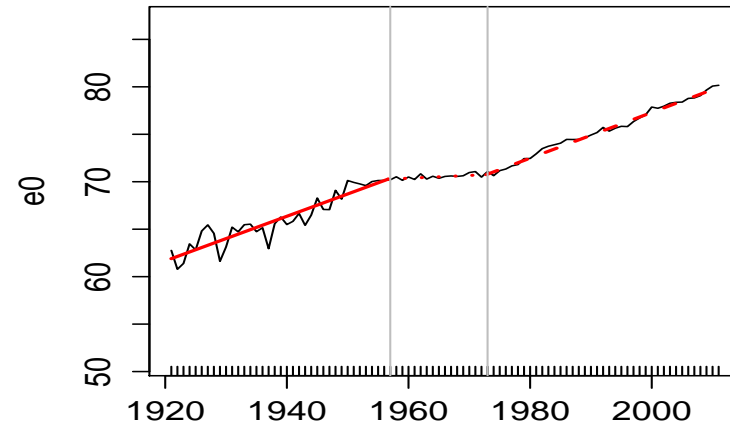


Breakpoints in Canadian e_x

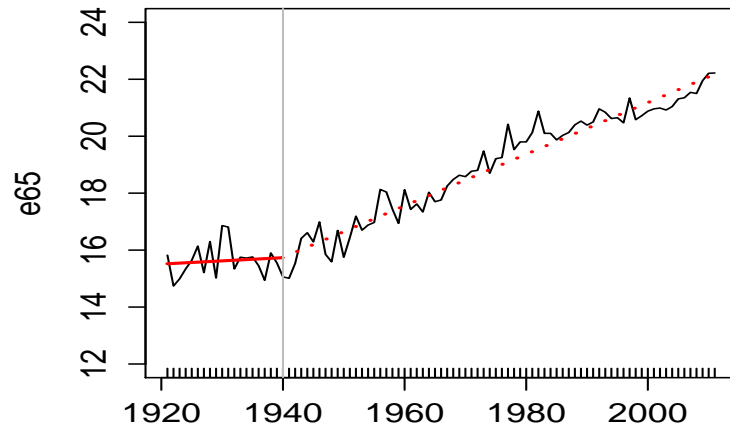
Females



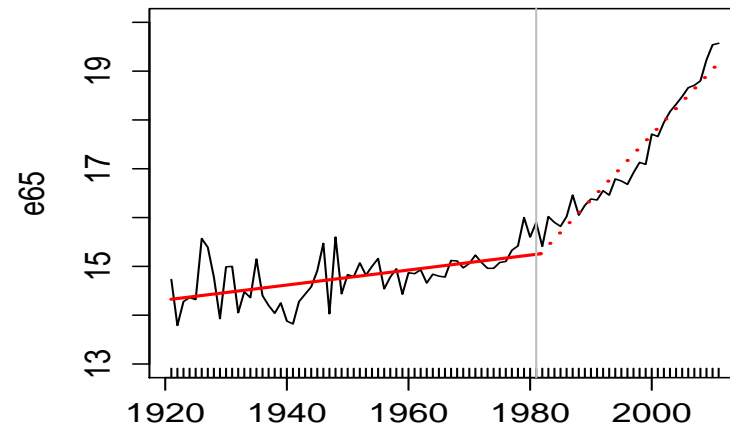
Males



Females



Males



The Model

Time-dependent GEV model to annual maximum provincial e_x :

$$GEV(\mu_t, \sigma_t, \xi_t) \quad \text{with} \quad \mu_t = \beta_0 + \beta_1 t; \quad \sigma_t = \sigma; \quad \xi_t = \xi$$

$$G(z_t) = \exp\left\{ - \left[1 + \xi \left(\frac{z - (\beta_0 + \beta_1 t)}{\sigma} \right) \right]^{\frac{-1}{\xi}} \right\}$$

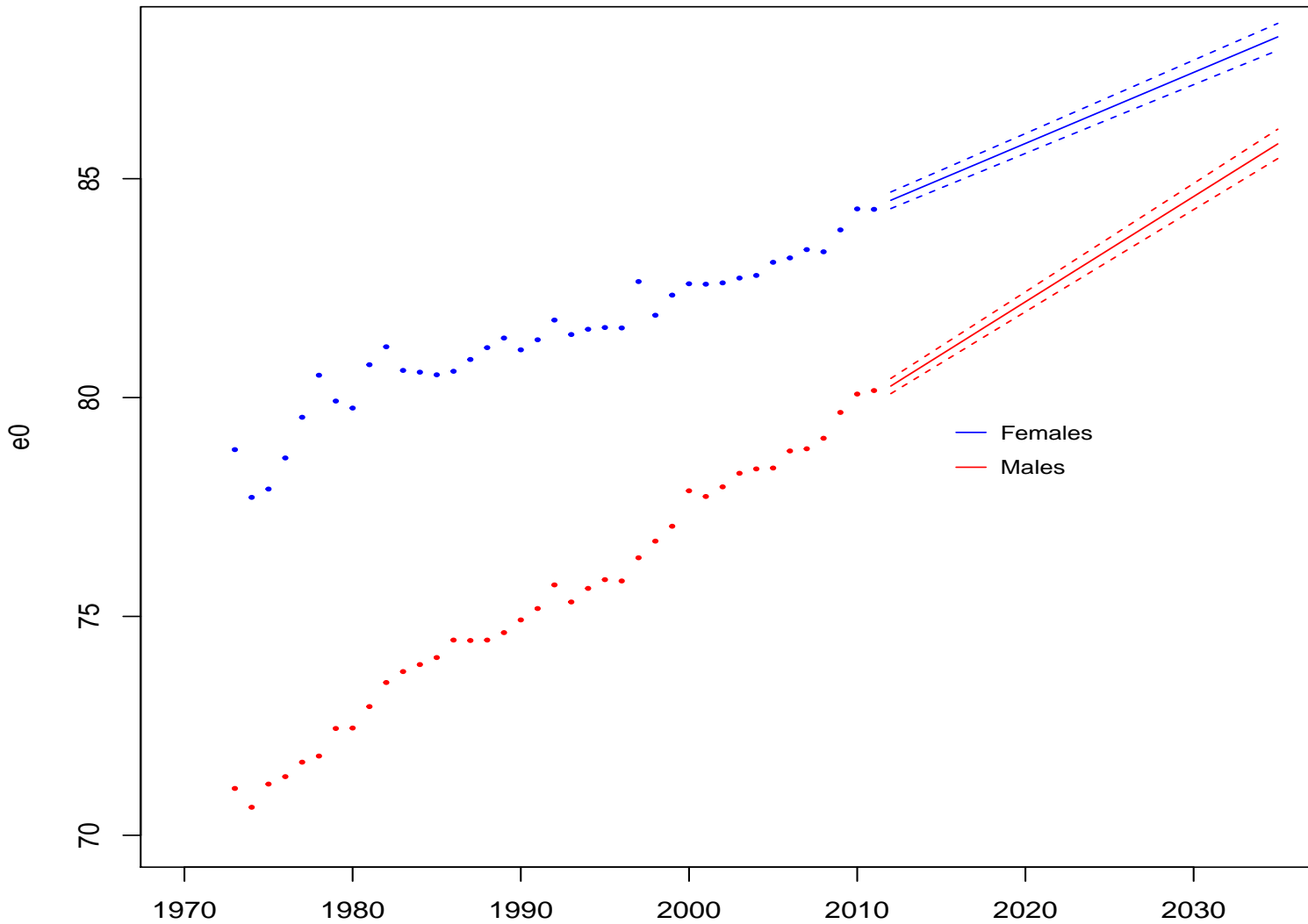
- ▶ Other forms of time dependence possible but linearity in μ reasonable and parsimonious choice.
- ▶ Parameters estimated jointly using maximum likelihood

Parameter estimates - Canada

	Neg. Likelihood	$\hat{\beta}_0$	$\hat{\beta}_1$	$\hat{\sigma}$	$\hat{\xi}$
Female e_0	30.8	76.4(0.11)	0.16 (0.003)	0.37 (0.030)	0.10
Male e_0	3.9	70.6 (0.10)	0.24 (0.004)	0.27 (0.03)	-0.34 (0.15)
Female e_{65}	45.2	15.4 (0.11)	0.09 (0.002)	0.42 (0.040)	-0.15 (0.08)
Male e_{65}	-3.60	15.1 (0.11)	0.13 (0.006)	0.27 (0.04)	-0.29 (0.14)

Table: Maximized negative log-likelihoods, parameter estimates and standard errors (in parentheses) of the Block Maxima Model

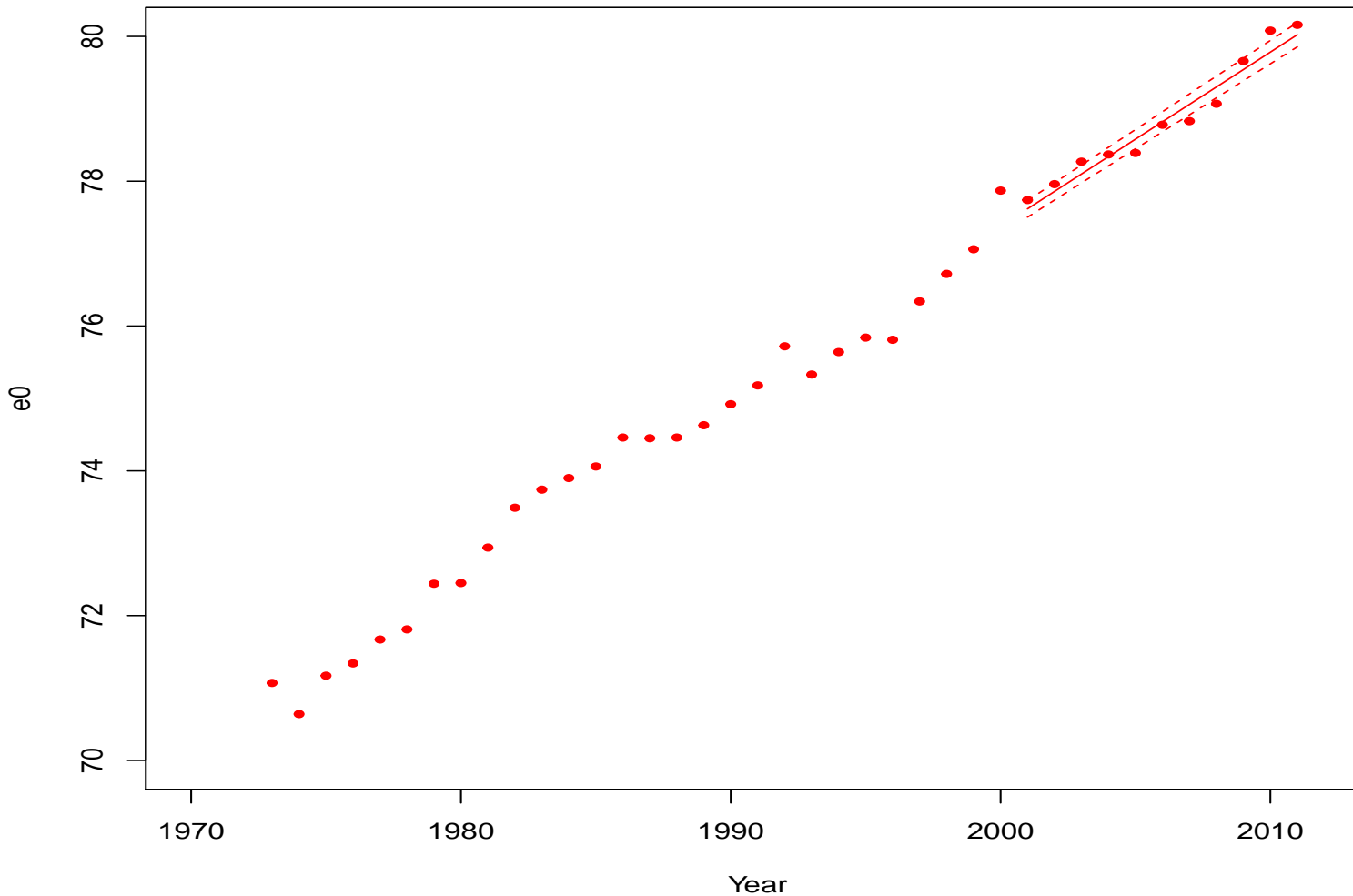
Projected e_0 - Canada



Probability statements - Canada

Year	$P(e_{0,f}^{max} > 87.5)$	$P(e_{0,f}^{max} > 89)$
2030	0.44	0.02
2035	0.99	0.11

Forecast Performance e.g. Canadian Males e_0



Comments

- ▶ Classical theory assumes that life expectancies between subregions are independent but there are dependencies
- ▶ Should have "large enough" number of subregions
- ▶ Data wastage?

Ongoing work: added sophistication

- ▶ Relax linearity assumption for time dependence to allow any time-varying shape for GEV parameters
- ▶ More flexible Dynamic Linear Model for time-varying parameters for forecasting
- ▶ Flexible GLM type framework for modelling
 - ▶ Vector Generalized Linear Models (Yee & Hastie, 2003), or
 - ▶ Generalized Additive Models for Location, Scale and Shape (Rigby & Stasinopoulos, 2001, 2005)

Take Aways

- ▶ Method can be used to make inferences about future maximum life expectancy for a region e.g a country using sub-regional information only e.g. states/ provinces/ prefectures
- ▶ Since a probability distribution is fitted, it is straightforward to obtain probabilities
- ▶ Ancillary benefit: if provinces with maxima also have high proportion of population then projecting median gives a workable estimate of overall country e_x
- ▶ Underlying theoretical model assumptions may be hard to achieve in practice but acid test is usually good assessment of empirical fit

References

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