

Assessing longevity inequality in the U.S.: What can be said about the future?

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Outline of the talk

1 Background

2 Reconciling mortality rates

3 Empirical results

- ▶ Case study on the “oldest-old”
- ▶ Projection into 2027

4 Conclusions

Some revealing facts about longevity

According to *The Guardian*:

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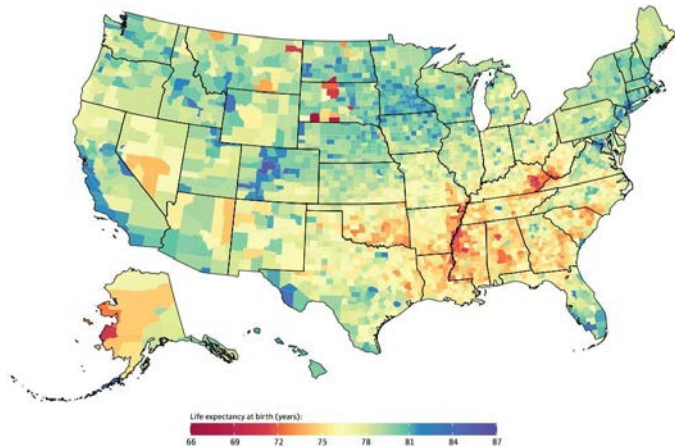
- Certain affluent counties in central Colorado: 87 years
- Several counties of North and South Dakota: 66 years
- The gap is predicted to grow even wider in future!

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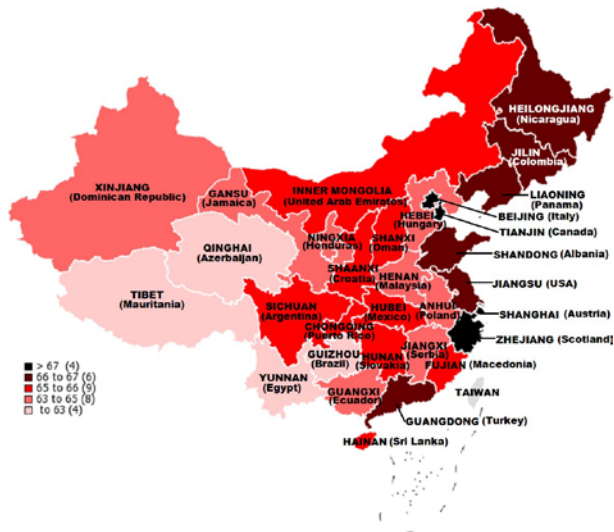
The big picture

Figure 1. Life Expectancy at Birth by County, 2014



Source: *JAMA Internal Medicine*

Something irrelevant (but interesting!)



Provincial HLE for males in 2015 with international matching



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How do we model mortality rates?

- **Central mortality rate** $m_{x,t}$, reflects the death probability for age x last birthday in the middle of the calendar year t and it is estimated by:

$$m_{x,t} = \frac{D_{x,t}}{E_{x,t}^c} .$$

- **Lee-Carter model (1992):**

$$\log(m_{x,t}) = a_x + b_x K_t \quad (1)$$

Forecast reconciliation

- Do we need coherent forecasts?
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Good question! We adopt a forecast reconciliation approach.



Forecast reconciliation: Baby steps

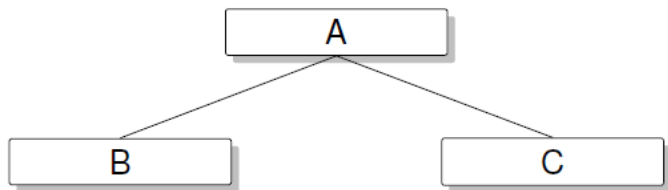


Figure: 2-level simple hierarchy

Forecast reconciliation: Baby steps

We first express the aggregation constraints in a matrix form:

- Define $\mathbf{y} = (a, b, c)$ as a vector that contains observations at all levels in the hierarchy;
- Define $\mathbf{b} = (b, c)$ as a vector that contains observations at the bottom level only.

The two vectors can then be linked by the equation

$$\mathbf{y}_t = \mathbf{S}\mathbf{b}_t, \quad (2)$$

where \mathbf{S} is a “**summing matrix**” of dimension 3×2 . It is given by

$$\mathbf{S} = \begin{pmatrix} 1 & 1 \\ 1 & 0 \\ 0 & 1 \end{pmatrix}. \quad (3)$$



Forecast reconciliation: Baby steps??

- Let $\hat{\mathbf{y}}_{T+h}$ denote the unreconciled h -step-ahead forecasts at all levels;
- Let $\tilde{\mathbf{y}}_{T+h}$ denote the correspondingly reconciled h -step-ahead forecasts which satisfy all aggregation constraints.

Any linear reconciliation method, according to Wickramasuriya *et al.* (2018), can be expressed as

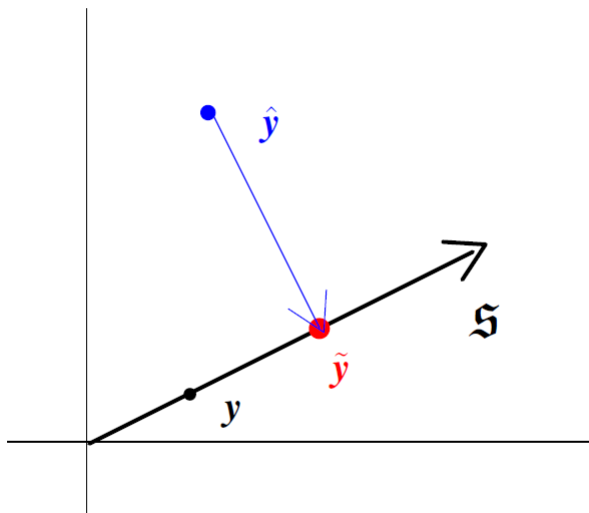
$$\tilde{\mathbf{y}}_{T+h} = \mathbf{S}\mathbf{P}\hat{\mathbf{y}}_{T+h}, \quad (4)$$

where in our example \mathbf{P} is a 2×3 matrix. The selection of \mathbf{P} is not unique and is a key step in forecast reconciliation. Wickramasuriya *et al.* (2018) propose a closed form estimation of the matrix \mathbf{P} by minimizing the trace of the covariance matrix of the in-sample *reconciled* forecast errors (MinT). The reconciliation matrix \mathbf{P} is given by

$$\mathbf{P} = (\mathbf{S}'\mathbf{W}_h^{-1}\mathbf{S})^{-1}\mathbf{S}'\mathbf{W}_h^{-1}, \quad (5)$$

where \mathbf{W}_h represents the covariance matrix of the h -step-ahead in-sample unreconciled forecast errors.

Forecast reconciliation: Understand it visually



Reconciling population exposure

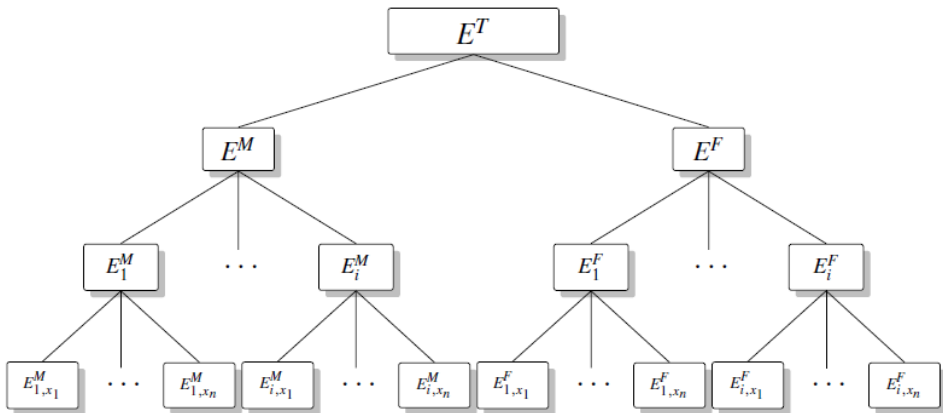


Figure: 4-level hierarchical tree for population exposure

Reconciling population exposure

For the hierarchical structure in the death counts, we have the following aggregation constraints at all times:

$$\sum_{x=x_1}^{x_2} E_{i_0,x}^j = E_{i_0}^j, \quad \forall i_0 \in [1, i], \quad j = M, F \quad (6)$$

$$\sum_{i_0=1}^i E_{i_0}^j = E^j, \quad j = M, F \quad (7)$$

$$E^M + E^F = E^T. \quad (8)$$

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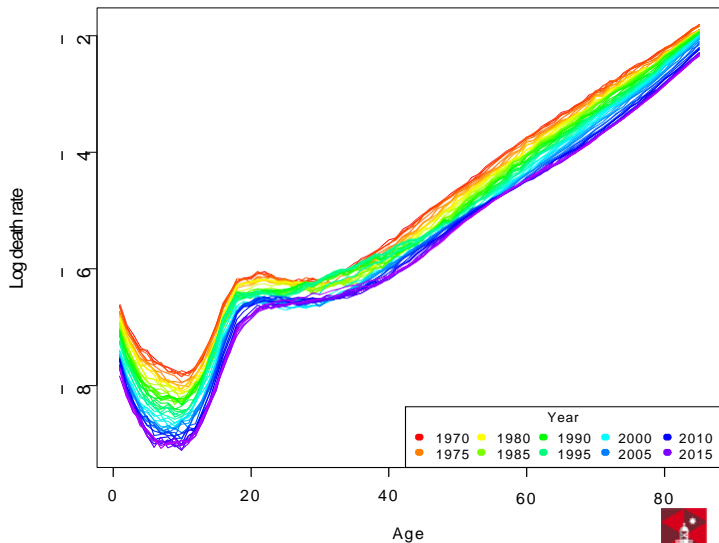
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Mortality forecasting

U.S.A.: male death rates (1970–2015)



Reconciling death counts

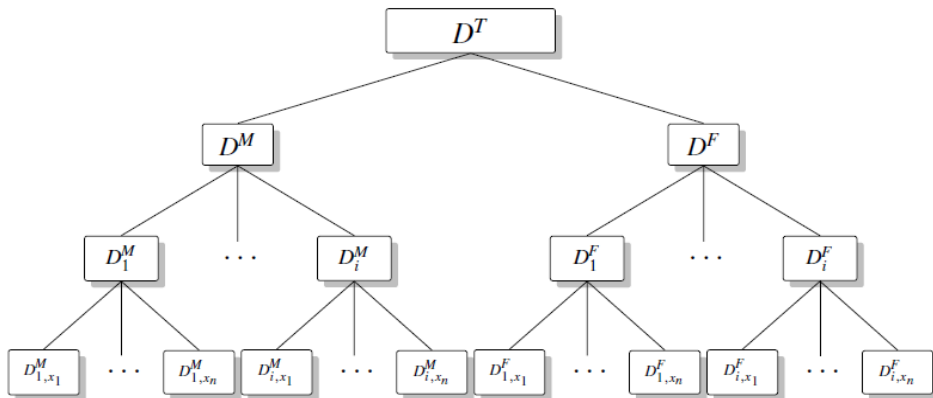


Figure: 4-level hierarchical tree for death counts

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Case study

The death and exposure data used to perform this case study are downloaded from the *CDC WONDER online database*.

- **Country:** the United States.
- **Number of states:** 50 plus District of Columbia.
- **Investigation period:** 1968–2017.
- **Age range:** 80+.

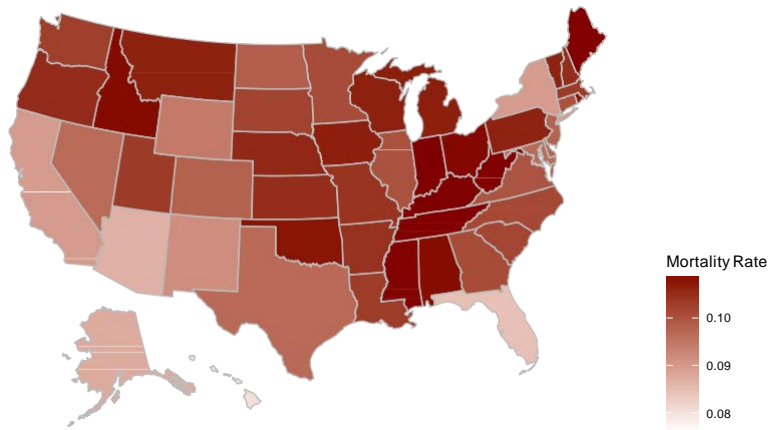


Mortality experience for selected U.S. states in 2017

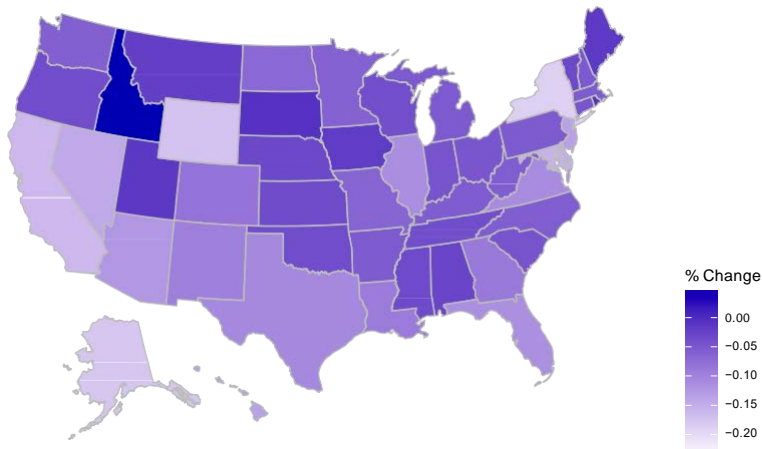
Rank	State	Crude rate	Rank	State	% change since 1990
1	District of Columbia	7.99%	1	District of Columbia	-24.19%
2	Hawaii	8.03%	2	New York	-18.69%
3	Florida	8.47%	3	Alaska	-17.58%
4	Arizona	8.68%	4	Wyoming	-17.13%
5	Alaska	8.76%	5	California	-16.03%
...
47	West Virginia	10.91%	47	Maine	-1.23%
48	Maine	10.91%	48	Utah	-1.02%
49	Tennessee	10.92%	49	Rhode Island	-0.39%
50	Kentucky	10.95%	50	South Dakota	-0.38%
51	Indiana	10.98%	51	Idaho	4.91%



U.S. oldest-old mortality rates in 2017



U.S. % change in oldest-old mortality rates: 1990–2017



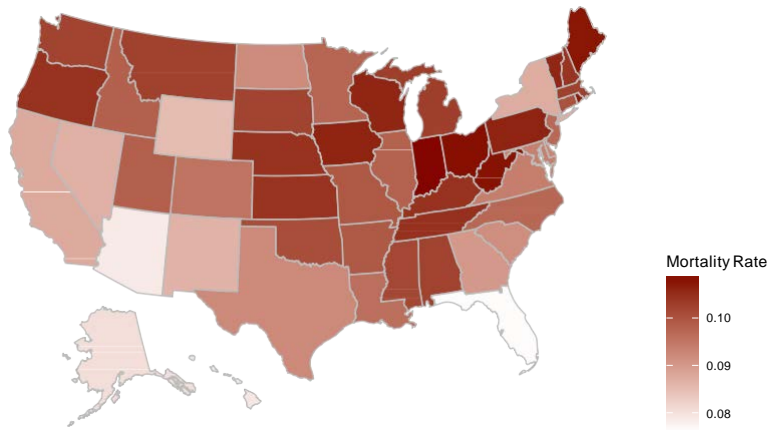
Mortality forecasts for selected U.S. states in 2027

Rank	State	Crude rate	Rank	State	% change since 2017
1	District of Columbia	7.38%	1	Georgia	-10.98%
2	Florida	7.64%	2	Nevada	-10.28%
3	Arizona	7.89%	3	South Carolina	-10.16%
4	Hawaii	7.93%	4	Wyoming	-9.97%
5	Alaska	8.06%	5	Florida	-9.74%
...
47	Iowa	10.65%	47	Ohio	-0.43%
48	Indiana	10.67%	48	Iowa	-0.37%
49	Rhode Island	10.72%	49	Pennsylvania	-0.16%
50	West Virginia	10.74%	50	Connecticut	-0.05%
51	Ohio	10.83%	51	South Dakota	0.06%

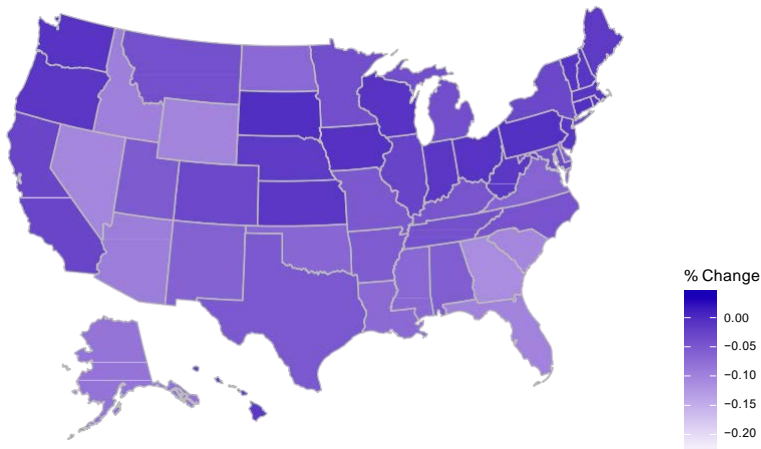
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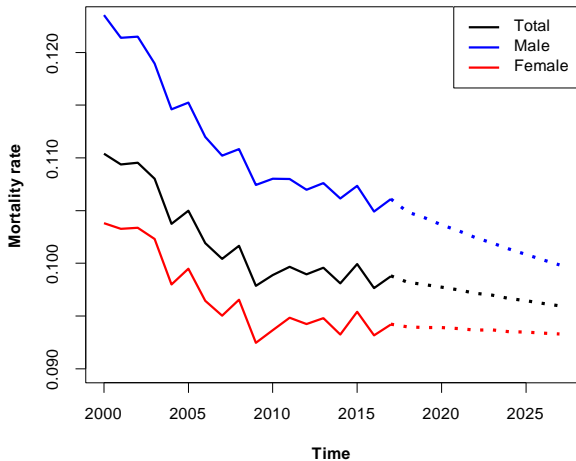
U.S. oldest-old mortality forecasts in 2027



U.S. % change in oldest-old mortality rates: 2017–2027



Oldest-old mortality rate forecasts in the U.S.: 2018–2027

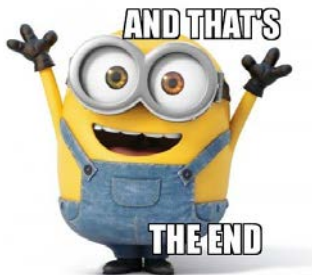


Findings

- The worst older-age mortality forecasts are generally observed in “rust belt states” such as Pennsylvania, West Virginia, Ohio and Indiana.
- Older-age males are predicted to experience a more rapid mortality improvement compared to females.
- Future older-age mortality improvement rate will tend to slow down across a majority number of states.



End of presentation



Thank you!

Any questions/ comments/ suggestions?

Contact email: han.li@mq.edu.au

Research Gate page: http://www.researchgate.net/profile/Han_Li51



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References

- Wickramasuriya, S. L., Athanasopoulos, G., and Hyndman, R. J. (2018). Optimal forecast reconciliation for hierarchical and grouped time series through trace minimization. *Journal of the American Statistical Association*, (just-accepted), 1–45.