

# Mortality collar

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# Programme

- 1 Introduction
- 2 The data and mortality model
- 3 The risk measure
- 4 The counterparts losses
- 5 The duality and results
- 6 Conclusion

# General context

## Motivation

- Some derivatives exist for hedging against mortality or longevity risks (Bauer, Börger & Rub 2008, Cairns, Blake & Dowd 2006, Cox, Lin & Milidonis n.d.),
- Some of them are on natural hedging (Cox & Lin 2004).

We develop here a contract based on natural hedging of downside risks transfer mechanism between a life insurer portfolio and an annuity's one with the mortality rate quantiles as bounds for a cap and a floor.

# The collar principle

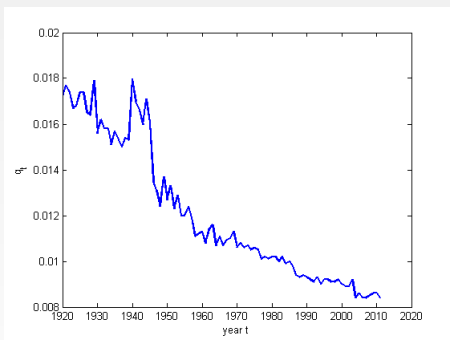
## a cap and a floor

At each constatation date :

- ✦ if  $q_{x,t} > q_{x,t}^{\beta}$  then the losses of the life insurer corresponding to the exceptional rising of the mortality from  $q_{x,t}^{\beta}$  to the observed  $q_{x,t}$  are compensated by the pension fund ;
- ✦ similarly, if  $q_{x,t} < q_{x,t}^{\alpha}$  then the losses of the pension fund are compensated by the life insurer ;
- ✦ at least if the mortality rate  $q_{x,t}$  lie in the collar no one have to pay to the other.

## The data

We consider French mortality data ( <sup>1</sup> ).



**FIGURE:** Historical mortality rate of French population (1920 to 2011)

# The mortality model

(Lee & Carter 1992) and (Cox et al. n.d.)

$$q_{x,t} = e^{\alpha_x + \beta_x \kappa_t + \epsilon_{x,t}} \quad (1)$$

## The mortality model

(Lee & Carter 1992) and (Cox et al. n.d.)

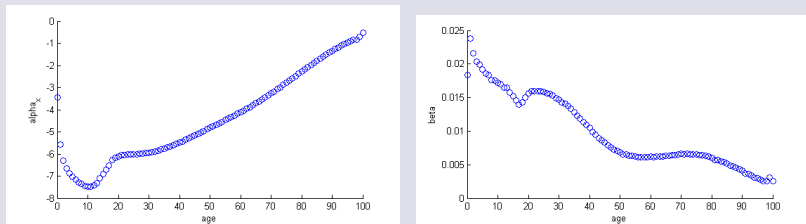


FIGURE: Estimated values of  $\alpha_x$  &  $\beta_x$

# The mortality model

(Lee & Carter 1992) and (Cox et al. n.d.)

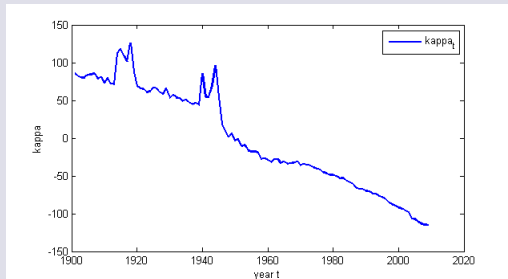


FIGURE: Estimated values of  $\kappa_t$



## The mortality model

(Lee & Carter 1992) and (Cox et al. n.d.)

TABLE: Statistics of error term from model 4

Parameter	Parameter value
Mean	$7.10^{-6}$
Standard deviation	10.5973
Kurtosis	10.8995
Skewness	0.1910

The kurtosis of  $v_t$  10.8995 and its skewness 0.1910.

# The mortality model

**TABLE:** Calibrated parameters of Regime-switching lognormal and lognormal models

Parameters	LN model	RSLN model
log-likelihood	<b>-411.4676568</b>	<b>-0.010717659</b>
$\mu$	$-3.66972 \cdot 10^{-6}$	
$\sigma$	10.54861	
$\mu_1$		0.032898375
$\mu_2$		-0.144669812
$\sigma_1$		8.504312799
$\sigma_2$		8.229352177
$\rho_{12}$		0.9999
$\rho_{21}$		$9.3982 \cdot 10^{-6}$

## The risk measure

### Utility function (Gerber & Pafumi 1998)

HARA-CRRA power utility function of the second kind :

$$u(x) = \frac{x^{1-c} - 1}{1-c} \text{ (with } c > 0\text{)}.$$

$$E_Q[V(X)] = \frac{E[V(X) u'(X)]}{E[u'(X)]} \quad (2)$$

With the choice of  $X = X_t(s) = q_{X_t,s}$  we have

$$E_Q[V(q_{X_t,s})] = \frac{E[V(q_{X_t,s}) q_{X_t,s}^{-c}]}{E[q_{X_t,s}^{-c}]} \quad (3)$$

## The life insurer's over quantile loss

### Projected liabilities differences

$$E_Q [V_1 (q_{x_t, s})] = E_Q \{ [LF (x_t, s) - VaR_\beta [LF (x_t, s)]] 1_{[LF > LF_\beta]} \} \quad (4)$$

# The pension fund over quantile loss

## Projected liabilities differences

$$E_{Q_t} [V_2 (q_{x_t, s})] = E_{Q_t} \{ [FP(x_t, s) - FP_\alpha(x_t, s)] 1_{[LF > LF_\alpha]} \} \quad (5)$$

# The pension fund over quantile loss

## Projected liabilities differences

$$EQ_t [V_2(q_{x_t, s})] = \frac{PN_{x_t} e^{Ri_{x_t} t}}{1 - e^{-Ri_{x_t}}} \left\{ e^{-Ri_{x_t} T_{x_t}^\alpha} \frac{E_t [q_{x_t, s}^{-c} 1_{[q < q^\alpha]}]}{E_t [q_{x_t, s}^{-c}]} - \frac{E_t [e^{-Ri_{x_t} T_{x_t}^\alpha} q_{x_t, s}^{-c} 1_{[q < q^\alpha]}]}{E_t [q_{x_t, s}^{-c}]} \right\} \quad (6)$$

# Duality

## Feasibility of the contract

The feasibility of the contract is conditioned by the fact that at the present time each party has the same expected loss :

$$E_{Q_t} [V_1 (q_{x_t,s})] = E_{Q_t} [V_2 (q_{x_t,s})] \quad (7)$$

. From this feasibility equation and with evident notations we get a relation between  $\alpha$  and  $\beta$  :

$$\begin{aligned} & \frac{C}{\bar{P}} (1 - e^{-Ri}) e^{Rs} \left\{ E_t [q^{1-c} 1_{[q > q^\beta]}] - q^\beta E_t [q^{-c} 1_{[q > q^\beta]}] \right\} \\ & = \left\{ e^{-RiT^\alpha} E_t [q^{-c} 1_{[q < q^\alpha]}] - E_t [e^{-RiT} q^{-c} 1_{[q < q^\alpha]}] \right\} \end{aligned} \quad (8)$$

# Duality

## The method

If  $\beta$  is known, we have to derive the right difference without knowing  $\alpha$  numerically by the least square of the two differences

$$(\sqrt{(E_{Q_t} [V_1(q_{x_t,s})] - E_{Q_t} [V_2(q_{x_t,s})])^2}).$$



## Conclusion

Remarks The solutions are functions







- 1 of the ratio of C and P to rebalance the payment if one counterpart pays more than the other,
- 2 but not of the length of the population, instead of the lengths are not the same.

The market data are the European interbank and zero-coupon bonds rates for the term structure and the inflation rates are from the Agence France Trésor consumer price index without tobacco.

# Conclusion

Discussion : work in progress...

- 1 The calibration of the CRRA coefficient is more "complex" than expected...,
- 2 simulations results, efficacy tests and continuous times interest and inflation rates models are coming soon.
- 3 since this contract deals with downside risks, what about extreme values theory ?

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